

A Survey of Recent Empirical Money Demand Studies

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This paper surveys a selected number of studies that evaluated the demand for money using the error-correction model approach in the 1990s across a range of industrial and developing countries. It briefly presents issues relevant to modeling and estimating the demand for money; and synthesizes information concerning variables, data period and frequency, unit root and cointegration techniques, stability tests, and findings in a tabular form. In addition, it presents estimated long-run income elasticity and elasticities or semi-elasticities for opportunity cost and other variables in a comparable framework. It aims to provide a reference tool for future research on demand for money in various countries. [JEL E41]

Demand for money plays a major role in macroeconomic analysis, especially in selecting appropriate monetary policy actions. Consequently, a steady stream of theoretical and empirical research has been carried out worldwide over the past several decades. The interest has, however, heightened in recent years, triggered primarily by the concern among central banks and researchers on the impact of the movement toward flexible exchange rate regime, globalization of capital markets, ongoing domestic financial liberalization and innovation, advancement in time series econometrics, and country-specific issues.

The extensive literature underscores two major points relevant to modeling and estimating the demand for money: variable selection and representation, and

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framework chosen. Failure to provide due consideration to these issues has tended to yield poor results. For the former, proper specification of opportunity cost variables happens to be the most important factor in getting meaningful results. Regarding the latter, the chosen system should be free of theoretical and estimation problems, and should perform well in empirical testing. The error-correction models (ECMs) have shown to meet these criteria.

This paper surveys a selected number of papers that applied the ECM approach to analyze the demand for money (of various definitions) during the 1990s in several industrial and developing countries.¹ The objective is to extract relevant information from these studies and provide it in a readily useable and comparable framework. In specific, the paper presents details concerning the techniques followed, variables chosen, periods and frequency selected, and major findings. In addition, it summarizes the long-run income elasticities, interest-rate semi-elasticities (or elasticities), and the coefficients of other relevant variables. It is hoped that the materials presented in this paper provide some reference points concerning the behavior of money demand in various countries, which in turn will help the policy makers in designing appropriate monetary policy actions and the researchers in carrying out further research.²

The paper is organized as follows: Section I briefly specifies the general framework that usually underlies the empirical formulation in estimating the demand for money. Section II carries out relevant discussion regarding the variables and estimation techniques, and summarizes information concerning various studies including the findings and estimated coefficients. Finally, Section III presents the conclusions.

I. General Framework

There is a diverse spectrum of money demand theories emphasizing the transactions, speculative, precautionary or utility considerations.³ These theories implicitly address a broad range of hypotheses. One significant aspect, however, is that they share common important elements (variables) among almost all of them. In general, they bring forth relationship between the quantity of money demanded and a set of few important economic variables linking money to the real sector of the economy (see Judd and Scadding, 1982, p. 993). What sets apart among these theories is that although they consider similar variables to explain the demand for money, they frequently differ in the specific role assigned to each. Consequently one consensus that emerges from the literature is that the empirical work is motivated by a blend of theories.

The general specification begins with the following functional relationship for the long-term demand for money:

¹This paper is based on Sriram (1999b, 1999c, and 2000). There have been other survey papers (for example, Judd and Scadding, 1982; Goldfeld and Sichel, 1990; Boughton, 1992; Laidler, 1993); but none of them focused exclusively on ECMs and covering a wide range of both industrial and developing countries.

²Refer to Ericsson (1998) for general issues concerning the empirical modeling of money demand.

³See, Laidler (1993) and Sriram (1999c), among others, for a survey of these approaches.

$$\frac{M}{P} = f(S, OC) \quad (1)$$

where the demand for real balances M/P is a function of the chosen scale variable (S) to represent the economic activity and the opportunity cost of holding money (OC). M stands for the selected monetary aggregate in nominal term and P for the price. Like in theoretical models, the empirical models generally specify the money demand as a function of real balances (see Laidler, 1993).⁴

II. Discussion on Variables and Estimation Techniques

Given the above general framework, this section provides a brief overview of issues concerning selection and representation of variables, modeling, and estimation. Sriram (1999c) presents detailed account of these issues, including relevant references justifying various approaches undertaken by the researchers. The literature shows that money demand has been estimated for various aggregates, their components, or certain combination of these components. As definitions of money differ across countries (see Boughton, 1992, and Kumah, 1989), measures considered, including divisia aggregates, also varied across studies. Scale variable is used in the estimation as a measure of transactions relating to the economic activity. It is usually represented by variables expressing income, expenditure, or wealth concept (although a host of other variables is discussed in the literature). The price variable is selected to follow closely the chosen scale variable, although consumer price index is the most commonly used measure.

One of the most important aspects of modeling the demand for money is the selection of appropriate opportunity cost variables. The literature has shown that studies which paid inadequate attention on this matter produced poor results. There are two major ingredients: (i) own-rate and (ii) alternative return on money. The former happens to be very important, especially if the financial innovation has been taking place in an economy (see Ericsson, 1998). The latter involves yields on domestic financial and real assets for a closed economy, and additionally on foreign assets for an open economy. A number of instruments are available to represent the yields on domestic financial assets. The yield on real assets is usually proxied by the expected inflation. And, on foreign assets by foreign interest rate or some form of exchange rate variable. Prior to selecting appropriate opportunity cost variables, careful attention should be paid on evaluating macroeconomic situation and developments in the financial system (including institutional details and the regulatory environment), and degree of openness of the economy.

The economic theory provides some guidance in reference to the relationship between demand for money and its arguments. As the scale variable represents the transactions or wealth effects, it is positively related to the demand for money. The

⁴Using the real money balance as the dependent variable will also mean that price homogeneity is explicitly imposed into the model. Additionally, there are less severe econometric problems associated with using real rather than nominal balances as the dependent variable (see Boughton, 1981, and Johansen, 1992b). And, majority of the empirical work does find evidence for the demand being for *real* balances.

own-rate is expected to be positively related as higher the return on money, less the incentive to hold assets alternative for money. Conversely, higher the returns on alternative assets, lower the incentive to hold money, and hence, the coefficients of alternative returns expected to be negative. The expected inflation generally affects the demand for money negatively as agents prefer to hold real assets as hedges during the periods of rising inflation. The foreign interest rates are expected to exert negative influence as increase in foreign interest rates potentially induce the domestic residents to increase their holdings of foreign assets which will be financed by drawing down domestic money holdings. Similarly, the expected exchange depreciation will also have a negative relationship. An increase in expected depreciation implies that the expected returns from holding foreign money increases, and hence, agents would substitute the domestic currency for foreign currency.⁵

The economic theory does not provide any rationale as to the correct mathematical form of the money demand function. There is consensus, however, that the log-linear version is the most appropriate functional form (see Zarembka, 1968). While money and scale variables typically enter in logarithms, interest rate variables appear either in levels or in logarithms. Consequently, estimates of the coefficient for the scale variable directly provides the measure of income elasticity, and those of interest rates show either elasticities or semi-elasticities depending on the way they are introduced in the formulation.

The partial adjustment framework was extremely popular in the 1970s. However, it was shown to suffer from specification problem and highly restrictive dynamics (see, for example, Cooley and LeRoy, 1981; Goodfriend, 1985; Hendry, 1979 and 1985; Hendry and Mizon, 1978). To counter these problems, two major solutions were proposed—modifying the theoretical base and improving the dynamic structure. The former led to buffer-stock models (BSMs), which were built upon the theory of precautionary demand for money (see, for example, Laidler, 1984; Cuthbertson and Taylor, 1987; Milbourne, 1988), and the latter to ECMs.⁶ The BSMs also ran into criticism, especially in their relevance in the empirical estimation (see Milbourne, 1988). Meanwhile, ECMs seem to be promising. An important aspect of these models is that the data characteristics are thoroughly examined before selecting the appropriate estimation techniques. Furthermore, lag structures are selected based on the data generating process of the economic variables and not on *a priori* based on the economic theory or naive dynamic theory.

The ECM is shown to contain information on both the short- and long-run properties of the model with disequilibrium as a process of adjustment to the long-run equilibrium. Granger (1983 and 1986) has demonstrated that the concept of stable long-run equilibrium is the statistical equivalence of cointegration. When cointegration holds and if there is any shock that causes disequilibrium, there exists a well-defined short-term dynamic adjustment process such as the error-

⁵Refer to Jusoh (1987) and Tan (1997) for reasons to expect positive relationship for expected inflation and expected exchange rate depreciation with the demand for real money respectively.

⁶In fact, Hendry, Pagan, and Sargan (1984) showed that PAMs and BSMs form the special cases of ECMs.

correction mechanism that will push back the system toward the long-run equilibrium. In fact, cointegration does imply the existence of a dynamic error-correction form relating to variables in question (see Engle and Granger, 1987). The major advantage of the error-correction modeling is that the economic theory is allowed to specify the long-run equilibrium while the short-run dynamics be defined from the data.

The earlier ECMs on money demand tended to be based on the single equation cointegrating relationship between money and the chosen scale variables as developed by Engle and Granger (1987). However, further research suggested that multivariate cointegrating vectors encompassing a broader number of variables provided a fuller characterization of the long-run determinants of demand. The specification of such multiple cointegrating vectors between nonstationary variables primarily employs the procedures developed by Johansen (1988) and Johansen and Juselius (1990) which make the original Engle-Granger framework a special case. However, as can be seen from Table 1, a number of other measures available to conduct the cointegration analysis.⁷

Table 1 also presents details relevant to modeling and estimating the demand for money from various studies. In specific, it summarizes information for a cross-section of developing and industrial countries, on monetary aggregates (nominal or real), scale variable(s), and the opportunity cost and other variables included; data period and frequency chosen; unit root, cointegration, and stability tests applied; nature of various time series (such as the order of integration and whether seasonally adjusted or not). It also presents the findings. The presentation of information will enable the researchers to draw some insights into the justification of selecting diverse set of variables and approaches across various countries.

Table 2 summarizes the long-run income elasticities and the semi-elasticities or elasticities of opportunity cost and other variables from those studies listed in Table 1. As the short-run dynamics can be potentially complicated, the table concentrates only on the long-run results. In order to promote comparability, the results are shown only for those studies which reported the long-term relationship (existence of cointegration). If more than one cointegration relationship is found, results are reported only for the preferred cointegration vector(s) as identified by the author(s), which not only meet a battery of statistical tests but also economically make sense with correct signs of the variables and meaningful size of coefficients.

Figures 1–3 show the distribution of income elasticities for real money as presented in Table 2 for components of narrow money, narrow money, and broad money respectively. The relevant descriptive statistics is shown in Table 3. It is clear from the table, the medians for all three groups are closer to one than to 0.5 thereby indicating that money does not play the role of transaction measure alone. There is no clear guidance from the theory or empirical studies regarding the acceptable magnitude on elasticities or semi-elasticities of the opportunity cost variables. The most relevant information will be the signs of the coefficients—positive for own-rate and negative for alternative return on money and expected

⁷Refer to Sriram (1999c) for a longer list of studies that applied the ECM framework to analyze the demand for money in the past two decades.

Table 1. Summary of Demand for Money Studies Involving Cointegration/Error-Correction Modeling in Selected Industrial and Developing Countries

Country/ Author(s)	Sample/ Period/ Frequency	Monetary Aggregate(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
Industrial countries											
Australia Lim (1993)	1977:4-1990:2 Quarterly	Real currency; ¹ real bank deposits; ¹ real nonbank deposits ¹ [GDPD-based]	Real GDP ¹	90-day bank bill rate; 2- and 5- year T-bond rate	Inflation rate (GDPD- based); structural dummy	ADF; P (1987)	90-day bank bill rate is I(0); others are I(1)	PH (1990) “fully modified” regression; ² JJ (1990); PO (1990)	...	Yes	Cointegrating relationships exist for both monthly and quarterly models for each money variable (without the 90-day bank bill rate); ECM shows some evidence for the significance of the 90-day bank bill rate in influencing the short-run of the monetary aggregates.
Canada Haug and Lucas (1996)	1953:1-1990:4 1968:1-1990:4 Quarterly	ln (real M1); ¹ ln (real M2); ¹ ln (real M2+); [GDPD-based]	ln (real GDP) [GDPD- based]	ln (91-day T-bill rate); ln (10-year T-bond rate)	...	DF	I(1)	AEG; DOLS; JJ (1990); PO (1990)	Hansen (1992)	No	Results vary depending on the cointegration tests selected and the combination of money and interest rates; however, stable long-term relationship is found among real M1, real GDP, and the 91-day T-bill rate.
Germany Deutsche Bundesbank (1995)	1970:1-1994:4 Quarterly	Log (M3/ GNPD) [M3 is adjusted for statistical breaks]	Log (real GNP) [GNPD-based]	Yield on domestic bearer debt securities outstanding (f); r-if ²	Seasonal dummies	ADF	I(1)	EG (1987)	...	Yes	Cointegrating relationship exists among money, interest rate r, and real GNP. The EC term is calculated as the avg. of previous four quarters, and has the negative coefficient which is significant.

Table 1. (continued)

Country/ Author(s)	Sample/ Period/ Frequency	Monetary Aggregate(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
Greece	1976:2-1994:4	In (M3:CP)	In GDP at factor cost in constant 1970 prices)	Net return on TD; interest rate spreads for repos and deposits; ³ LIBOR	DEPR using NEER; inflation rate; seasonal and structural dummies	ADF	I(1)	EG (1987); J (1988); J (1991a); J (1992a); J (1992b)	Chow	Yes [General to Specific Approach]	Cointegrating relationship among money, scale variable, inflation rate, and domestic interest rates and the spreads; stable ECM.
Erisson and Sharma (1998)	Quarterly										
Italy	1963:1-1987:4	Log (M2/GDP) ¹	Log (real GDP) ¹	R = alternative return on M2 minus own-rate ^d	Variables to express learning curves after the introduction of BOTs and CCTs ⁵	ADF; PO (1990); PP (1988)	I(1)	EG (1987)	Chow	Yes [General to Specific Approach]	Cointegration relationship can be obtained only after the addition of learning curve variables. Demand for M2 is significantly affected by the introduction of new financial instruments.
Muscetti and Papi (1990)	Quarterly										
Japan	1973:1-1988:4	In (real M2) ¹	In (real GNP) ¹	In (1+R) ^e	In (real XR); ¹ inflation rate; ¹ ln (GNPD) ¹	DF; ADF; PP (1988)	I(1)	AEG	Ashley (1984); Chow; CUSUM; CUSUMSQ	Yes	Cointegrating relationship among real GNP, real wealth, and real XR; stable ECM throughout the sample period.
Arize and Shwiff (1993)	Quarterly										
New Zealand	1965:2-1989:4	Log (M3)	Log (real GDP)	Annual rate on S-T trading bank loans	Log (GDPD)	DF	I(1)	J (1988); JJ (1990)	...	No	Cointegration without interest rate for the sub sample; and with interest rate for full sample.
Orden and Fisher (1993)	1965:2-1984:2	Quarterly									

Table 1. (continued)

Country/ Author(s)	Sample/ Period/ Frequency	Monetary Aggregates(s)	Scale variable(s)	Determinants		Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
				Interest rate(s)	Other(s)						
Norway Bårdsen (1992)	1967:3-1989:4 Quarterly	In (NMX)	ln (real GDE)	Interest rate on DD and TD; yield on long-term private bond; 3-month euro-krone rate	ln (GDED)	not explicitly shown	I(1) except for 3-month euro-krone rate (which may be stationary around a trend)	J (1988); JJ (1990)	Chow	Yes	At least two and possibly up to five cointegration vectors exist; money is endogenously determined by prices, real expenditure, and interest rates.
Switzerland Chowdhury (1995)	1973:2- 1991:4 Quarterly	Log (real B) Log (real MI)	Log (real GDP)	S-T (3-month TDR on Euro deposits in banks rate Swiss francs); L-T return on federal bonds)	NEER; London clearing banks rate PP (1988)	ADF; KPSS (1992); PP (1988)	I(1)	J (1988); JJ (1990)	Chow	No	Demonstrates the importance of including variables expressing foreign influence in an open economy; without adding exchange rate no cointegration is found.
United Kingdom Drake and Chrystal (1994)	1976:2-1990:3 Quarterly	In (M1d); In (M2d); In (M3d) where d stands for divisia aggregates	ln (real GDP)	Benchmark rate of interest; own rates of interest on M2d and M3d	ln (GDPD); inflation [GDPD; based]; implicit divisia rental price or user cost indices for M1d, M2d, and M3d; dummy variable	DF; ADF; PP (1988)	I(1) except for implicit divisia rental price or user cost indices for M1d and M3d which are I(0)	J (1988); JJ (1990)	Chow; CUSUM; CUSUMSQ	Yes [General to Specific Approach]	Company sector money demand; cointegrating relationship exists for all monetary aggregates. ECMs indicate that the speed of adjustment of the EC term is faster for M1d than for M2d and M3d.

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregates(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
United States Miller (1991)	1959:1-1987:4 Quarterly	In (adjusted B); In (M1); In (M2); In (M3)	In (real GNP) CPR; In (dividend- price ratio)	In (4-6 month CPR); In (real GNP)	I(1)	DF; ADF	EG; AEG	...	Yes	Cointegration relationship exists among M2, real GNP, IPD, and the CPR. ECM for M2 suggests valid and significant error-correction term.	
Baba, Hendry, and Starr (1992)	1960:3-1988:3 Quarterly	In (M1/ IGNPD) ¹	In (real GNP) ¹	In (real GNP) ¹	J (1988)	I(1)	J (1988); JJ (1990)	Chow	Yes	Stable cointegrating demand function for real M1 (with the arguments which include inflation, real income, long-term bond yield and risk, T-bill interest rate, and learning curve weighted yields on newly introduced instruments in M1 and non-transactions M2).	
McNown and Wallace (1992)	1973:2-1988:4 Quarterly	Log (real M1); Log (real M2)	Nominal T-bill rate	Log (NEER)	ADF	I(1)	J (1988); JJ (1990)	...	No	[General to Specific Approach]	
Mehra (1993)	1953:1-1991:2 Quarterly	In (M2/IGNPD)	In (real GNP)	In (R-RM2) ² ...	ADF	Interest rate is I(0); others I(1)	OLS; IVT	Chow	Yes [OLS and IVT]	Example of a model that estimates both the long- and short-run coefficients in one step.	
											Cointegrating relationship for real M2 and real GNP; money demand function is stable throughout the sample period.

A SURVEY OF RECENT EMPIRICAL MONEY DEMAND STUDIES

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregates(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
Developing countries											
<i>Argentina</i>	1935:1-1962:4	ln (M1/WPD); ln (M2/WPI)	ln (real NNI)	...	Inflation rate [WPI-based]	ADF	I(1)	J (1988); JJ (1990)	...	Yes	Cointegration relationship exists among real money (M1 and M2), real NNI and the inflation rate. ECM finds relationship between real money and inflation.
<i>Choudhry (1995)</i>	1946:1-1962:4 Quarterly										
<i>Bolivia</i>	1980:9-1988:12	ln (B/CPD); ln (M1/CPI); ln (M2/CPI)	Expected inflation; inflation uncertainty	ADF	I(1)	J (1988); JJ (1990)	...	Yes	The null hypothesis of at least one cointegrating vector is not rejected. ECM contains time- varying EC term, estimated by Kalman filtering technique.
<i>Hanahan, and McNeilis (1993)</i>	Monthly										
<i>Cameroon</i>	Fielding (1994)	1976:1-1987:2 Quarterly	ln (BM/CPI)	ln (real GDP adjusted for terms of trade)	ln (1+CBDR) navart; quarterly dummy variables ⁸	DF; Hylleberg and others (1990)	I(1)	JJ (1990)	Chow [for ECM]	Yes	Three cointegrating relationships among real BM, real GDP, inflation, interest rate and navart. ECM passes diagnostic tests; EC term has a nearly unit coefficient.
<i>China</i>	Hafer and Kutan (1994)	1952-88 Annual	Log (currency); Log (currency plus SD)	Log (N/IRPD); Log (N/NID)	Log (one-year interest rate on SD)	DF	I(1)	J (1988); JJ (1990)	Cointegrating relationship exists only when NID (and not RPD) is used as a price variable;
											plus SD is the preferred measure of the monetary aggregate.

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregate(s)	Scale variable(s)	Determinants			Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
				Interest rate(s)	Other(s)	Unit Root Test(s)				
Tseng and others (1994)	1983:1-1988:4 1989:1-1993:4 1983:1-1993:4 Quarterly	ln (CC/RP) ¹ ln (M1/RP) ¹ ln (M2/RP) ¹ Interest rates exert significant influence on M1 and M2 in the 1989:1-1993:4 subperiod.	Real interest rate for the M1 and M2 equations for 1989:1-1988:4 [for 1983:1-1988:4]	Quarterly inflation rate (RP1-based)	ADF	I(1)	EG; JJ (1988; JJ (1990)	Chow	Yes	All monetary aggregates are sensitive to inflation although its impact drops during the 1989:1-1993:4 subperiod.
<i>Côte d'Ivoire</i> Fielding (1994)	1974:3-1987:4 Quarterly	ln (BM/CP)	ln (real GDP adjusted for terms of trade)	ln (1+CBDR)	In (1+ π); mavart; quarterly dummy variables ⁸	DF; Helleberg and others (1990)	I(1)	JJ (1990)	Chow	Yes
<i>India</i> Moosa (1992)	1972:1-1990:4 Quarterly	Log (CC/CPI); Log (NM/CPI); Log (BM/NM plus QM)/CPI	Log (MMR; rate offered in Bombay interbank market)	...	DF; ADF	I(1)	EG; AEG; CRDW; J (1988; JJ (1990)	...	Yes	Cointegration relationship exists for real money (except for BM using AEG) with IO and MMR. More stable relationship for CC and NM than for BM. ECMS show better results for CC and NM than for BM.

A SURVEY OF RECENT EMPIRICAL MONEY DEMAND STUDIES

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregates(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
<i>Indonesia</i> Price and Insukindro (1994)	1969:1-1987:4; Quarterly	ln (real CHP); ln (real DD)	ln (real GDP)	Rate of return on TD and on SD; LIBOR	Dummy variable for 1983 [for ECM]	DF; ADF	I(1)	EG; J (1988)	Chow; Salklever (1976) dummy approach [for ECM]	Yes	EG: weak evidence of cointegration relationship for currency; J (1988) finds up to 2 cointegrating vectors for both money equations. ECM does not find LIBOR being an important variable.
Deakle and Pradhan (1997)	1974:95 Annual	Log (NM); Log (BM); Log (real NM); Log (real BM)	Log (real GDP)	TDR [for NM]; MMR- TDR weighted by the share of QM in BM;	Log (CPI)	ADF	I(1) except for Log (CPI) which is I(0)	J (1988); JJ (1990)	...	No	No cointegrating relationship for any definition of money.
<i>Iran</i> Bahmani- Osokoee (1996)	1959-90 Annual	Log (M1/GDP); Log (GDP in 1980 prices)	Inflation; Log (official XR); Log (black market XR)	ADF; Peron (1989)	I(1)	J (1988); JJ (1990)	...	No	The most suitable model is the one that applies the black market XR with real GDP and inflation to explain demand for real M2.		
<i>Kenya</i> Adam (1992) among	1973:1-1989:2 Quarterly	Log (M0/CPD); Log (M1/CPD); Log (M2/CPD); Log (M3/CPD); Log (M3d/CPD)	Log (GNY/ r = quarterly yield on T-bill where M3d is divisia M3)	Expected DGP using parallel market XR; adjusted for changes in terms of trade	DF; CRDW	I(1)	J (1988); JJ (1990)	...	Yes	Two cointegrating vectors 5 variables for each monetary aggregate. ECMS validate the cointegrating relationships.	

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregate(s)	Scale variable(s)	Determinants	Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings	
Fielding (1994)	1975:2-1989:2 Quarterly	ln (BM/CPI)	ln (real GDP adjusted for terms of trade)	Interest rate(s) Other(s)	ln (1+T-bill rate) ln (1+DEPR) using parallel market XR; navarr; navarr; quarterly dummy variables ⁸	DF; Helleberg and others (1990)	I(0) for ln (1+DEPR); I(1) for others	JJ (1990)	Chow [for ECM]	Yes	Three cointegrating relationships among real money, real GDP, inflation, interest rate, navarr, and navarr. The EC term is calculated based on the residuals from the first two cointegrating vectors. S-T elasticities are smaller than those of long run.
Korea Arize (1994)	1973:1-1990:1 Quarterly (both)	ln (M1/CPI); ln (M2/CPI)	ln (real GDP)	CBR; interest rate	Expected rate of inflation; EER; standard deviation of the change in S-T interest rates in 9 industrial countries; uncovered interest rate differential in favor of foreign country	ADF; Helleberg and others (1990); TD on NCB; weighed avg. of S-T interest rates in 9 industrial countries; measures the change in circumstances	I(1)	EY (1987); J (1988); JJ (1990)	Chow	Yes	Two to three cointegrating vectors among real money M1 and M2), real income, interest rate, and foreign exchange rate risk and return. Well-specified ECM.
Lebanon Eken and others (1995)	1964-93 Annual	Log (B/CPI); Log (M1/CPI); Log FCDS; ¹⁰ Log (M2LL /CPI); ¹⁰	Log (real GDP); Log (U.S. dollar- denominated GDP)	... Log (CPI); Log (U.S. CPI); expected inflation; wa- year dummy	PP (1988)	I(1)	EG (1987); PO (1990)	Cointegrating relationship exists between various definitions of money and with real GDP, prices, and domestic inflation.	

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregate(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
<i>Malaysia</i> Sriram (1999a)	1973:8- 1995:12 Monthly	ln (M2/CPI)	ln (IP)	CBTDFM; discount rate on 3-month T-bills	Expected inflation; nominal XR; seasonal and structural dummies	DF; ADF	In (IP) and expected inflation are I(0); others are I(1)	J (1988); JJ (1990)	Chow	Yes [General to specific Approach]	Cointegration relationship exists between real M2 and its determinants under both the closed- and open-economy framework; fairly stable ECMs under both situations.
<i>Mexico</i> Khamsi and Leone (1999)	1983:1- 1997:6 Monthly	ln (CC/CPI)	ln (real private consumption expenditure)	60-day TDR	Inflation	ADF	I (1)	J (1988); JJ (1990)	Chow	Yes	Cointegration relationship among real CC, scale variable, and 60-day TDR; stable ECM.
<i>Morocco</i> Hoffman and Tahiri (1994)	1959:1-1988:2 Quarterly	Log (M1); Log (M2)	Log (GDP/ CPI); Log (GNP/ CPI)	Swiss S-T interest rate; interest rate on TD	Log (CPD); seasonal dummies	ADF; KPSS (1992)	I(1) possibly about a deterministic trend; KPSS test fails to reject the null of stationary for Swiss S-T interest rate adjusted for TDR	J (1988); J (1991b); JJ (1990); OLS; DOLS	Hansen and Johansen (1993)	No	Single cointegrating vector among measures of nominal money, prices, real income, and Swiss S-T interest rate.
<i>Nigeria</i> Fielding (1994)	1976:1-1989:2 Quarterly	ln (BM/CPI)	ln (real GDP adjusted for terms of trade)	ln (1+T-bill rate)	ln (1+ π_T); ln (1+DEPR) using parallel market XR; navarit; ⁸ seasonal dummies	DF; Hylleberg and others (1990)	I(0) for ln (1+DEPR); I(1) for others	JJ (1990)	Chow [for ECM]	Yes	One cointegrating relationship among real money, real GDP, inflation, interest rate and navarit.

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregates(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
Teriba (1997)	1960-94 Annual 1962:1-1995:2 for M1; and 1962:1-1992:4 for M2 Quarterly	Log (COB); Log (M1); Log (M2) for M1; and for M2 Quarterly	Log (real DA)	Log (interest rate for 12- month TD); Log (interest rate for 3- month TD)	Log (DAD); Log (LTBR in Nigeria/LTBR in the United States)	DF	I(1) except for Log M1 (I(2)) and for parallel market XR (I(0))	EG; AEG	...	Yes	Cointegration relationship exists among the monetary aggregates, DA, DAD, and interest rates. Foreign opportunity cost variable has influence on M1 equation only.
Pakistan Arize (1994)	1973:1-1990:1 Quarterly	ln (M1/CPD); ln (M2/CPD)	ln (real GDP) [WPI-based]	CMR; Govt. bond yield; weighted avg. of S-T interest rates in 9 industrial countries; uncovered interest rate differential in favor of foreign country	Expected rate of inflation; EER; standard deviation of the change in the log of the EER; dummy variable to measure the change in circumstances	ADF; Hylleberg and others (1990); Osborn (1990); Hasza and Fuller (1982); Peron (1988)	I(1)	EY (1987); J (1988); JJ (1990)	Chow	Yes	Two to three cointegrating vectors exist among real money (both M1 and M2), real GDP, interest rate, and foreign exchange rate risk and return. Well-specified ECM.
Hossain (1994)	1951-91 1972-91 Annual vectors	Log (M1/CPD); Log (M2/CPD)	ln (real GDP)	In yield on Govt. bonds; In (market call rate of interest)	Expected inflation	DF; ADF	Expected inflation is I(0); others	EG; AEG; CRDW; JJ (1990)	...	No	EG, AEG, and CRDW tests show conflicting results. But JJ (1990) test finds 2 cointegrating among money, real GDP, and call rate of interest for 1972-91 and one for 1953-91. M1 is found to be more stable than M2.

Table 1. (continued)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregates(s)	Determinants			Unit Root Test(s)	Order of Integration	Cointegration Technique(s) Test(s)	Stability Test(s)	Error- Correction Model (ECM)	Findings
			Scale variable(s)	Interest rate(s)	Other(s)						
Singapore Arze (1994)	1973:1-1990:1 Quarterly	ln (M1/CPD); ln (M2/CPD)	ln (real GDP) [WPI-based]	CMR; 3-month FDR; weighed avg.	Expected rate of inflation; EER; standard deviation of rates in 9 industrial countries; uncovered interest rate differential in favor of foreign country	ADF; Helleberg and others (1990); Osborn (1990); Hasza and Fuller (1982); Perron (1988)	I(1) except for expected rate of inflation which is I(0)	EY (1987); J (1988); JJ (1990)	Chow	Yes	2-3 cointegrating vectors among real money (both M1 and M2), real GDP, interest rate, and foreign exchange rate risk and return. Well-specified ECM.
Dekle and Pradhan (1997)	1975-95 Annual	Log (NM); Log (BMD); Log (real NM); Log (real BM)	Log (real GDP)	TDR [for NM]; MMR- TDR weighted by the share rate QM in BM; LIBOR	Log (CPI); expected depreciation rate	ADF	I(1)	J (1988); JJ (1990)	...	No	Cointegrating relationships for nominal NM and BM.
Thailand Dekle and Pradhan (1997)	1978-95 Annual	Log (NM); Log (BMD); Log (real NM); Log (real BM)	Log (real GDP)	TDR [for NM]; MMR- TDR weighted by the share of QM in BM;	Log (CPI)	ADF	I(1)	J (1988); JJ (1990)	...	No	Cointegrating relationship for nominal NM only.

Table 1. (concluded)

Country/ Author(s)	Sample Period/ Frequency	Monetary Aggregates(s)	Scale variable(s)	Determinants	Unit Root Test(s)	Order of Integration	Cointegration Technique(s)/ Test(s)	Stability Tests(s)	Error- Correction Model (ECM)	Error- Correction Model (ECM)	Findings
<i>Uruguay</i> Treichel (1997)	1963-95 Annual 1990-95 Monthly	In (M2/CPD); In (M4/CPD)	In (real GDP) on T-bill;	Monthly yield seasonal dummies rate; MMR	ADF I(1) except for inflation rate which is I(0)	I(1) except for inflation rate which is I(0)	AEG; J(1988); J(1990)	Recursive Chow [for ECM]	Yes	Stable long-term relationship among real money, real GDP, and the monthly yield on T-bill. Stable ECM.	

Note: The following abbreviations are used:

Monetary aggregates: B = base money; BM = broad money; CHP = currency held by public; CC = currency in circulation; COB = currency outside banks; DD = demand deposits; NM = narrow money; QM = quasi-money; SD = savings deposits; and TD = time deposits.

Scale variable: DA = domestic absorption; GDE = gross domestic expenditure deflator; GDP = gross national product; GNP = gross national product; IP = index of industrial production; IO = industrial output; NI = national income; and NNI = net national income.

Interest rate: CMR = call money rate; CBDR = Central Bank discount rate; CPR = commercial paper rate; CBR = corporate bond rate; FDR = fixed deposit rate; LIBOR = London interbank offered rate; LTBR = Long-term borrowing rate; MMR = money market rate; CBTID3M = Three-month deposit rates at commercial banks; TDR = time deposit rate; T-bill = Treasury bill; and T-bond = Treasury bond.

Exchange rate: DEPR = depreciation; XR = exchange rate; EER = effective exchange rate; and NEER = nominal effective exchange rate.

Prices: CPI = consumer price index; RPI = retail price index; and WPI = wholesale price index.

Deflators: DAD = domestic absorption deflator; GDED = gross domestic expenditure deflator; GDPD = gross domestic product deflator; GNPD = gross national product deflator; IGDPD = implicit GDP deflator; IGNPD = implicit GNP deflator; IPD = implicit price deflator; and NID = national income deflator.

Unit root tests: ADF = augmented Dickey-Fuller; CRDW = cointegration regression Durbin-Watson; DF = Dickey-Fuller; J (1988) = Johansen (1988); KPSS = Kwiatkowski, Phillips, Schmidt, and Shin (1992); P (1987) = Phillips (1987); PO (1990) = Phillips and Ouliaris (1990); and PP (1988) = Phillips and Perron (1988).

Cointegration tests: AEG = augmented Engle and Granger; CRDW = Cointegration regression Durbin-Watson; DOLS = dynamic ordinary least squares of Stock and Watson (1993); EG = Engle and Granger; EY = Engle and Yoo (1987); IVT = instrumental variable technique; J (n) = Johansen (n) where n stands for 1988, 1991a, 1991b, 1992a, 1992b respectively; JJ (1990) = Johansen and Juselius (1990); OLS = ordinary least squares; PH = Phillips and Hansen (1990); and PO (1990) = Phillips and Ouliaris (1990).

General: avg. = average; CB = corporate bonds; EC = error-correction; Govt. = Government; NCB = nationwide commercial banks; L-T = long-term; and S-T = short-term.

¹Seasonally adjusted.

²Where "if" stands for time deposit rate of deposits between DM 100,000 and DM 1 million.

³Spreads between yield on T-bill and net return on time deposits and between yield on T-bill and net return on repurchase agreements respectively.

⁴Own-rate is interest rate on bank deposits, net of taxes; and alternative return is yield on longer-term government debt.

⁵BOT stands for Buoni Ordinari del Tesoro and CCT for Certificati di Credito del Tesoro.

⁶R is defined as the three-month average Gensaki rate minus the average return on holding broad money defined as weighted average of the interest rate on three-month certificates of deposit and the guideline three-month deposit rate.

⁷R = own rate of return for M2 (weighted average of explicit interest rates paid on the components of M2) minus RM2 (four-six month CPR).

⁸mavar is annual moving average of changes in inflation calculated as $[\ln(p_t) - \ln(p_{t-1})]/t$, and mavar is for interest rates.

⁹Defined as one-year time deposit rate minus the rate of inflation.

¹⁰FCDS and M2LL stand for U.S. dollar-denominated deposits and Lebanese pound component of M2 respectively.

Table 2. Coefficients of Long-Run Demand for Money Estimated Under ECM Framework in Selected Countries¹

Study	Period/ Frequency	Method	Money ³	Elasticity		Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴
Industrial countries								
Canada Haug and Lucas (1996)	1953:1– 1990:4 Quarterly	DOLS	<i>mI</i>	0.420				-0.033*
Germany Deutsche Bundesbank (1995)	1970:1– 1994:4 Quarterly	EG (1987)	<i>m³</i>	1.400				-1.220
Greece Ericsson and Sharma (1998)	1976:2– 1994:4 Quarterly	J (1988); JJ (1990)	<i>m³</i>	1.220		7.650 & 7.020	-10.090	-3.380
Italy Muscatelli and Papi (1990)	1963:1– 1987:4 Quarterly	EG (1987)	<i>m²</i>	1.367			-2.082	-0.352 ⁵
Japan Arize and Shwiff (1993)	1973:1– 1988:4 Quarterly	AEG	<i>m²</i>	0.641 & 0.378 ⁶				0.094* ⁷
New Zealand Orden and Fisher (1993)	1965:2– 1989:4 Quarterly	J (1988); JJ (1990)	<i>M³</i> <i>M³</i>	0.410 0.630		1.130 1.020		-0.014 -0.001

Table 2. (continued)

Study	Period/ Frequency	Method	Money ³	Elasticity		Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴
Norway Bårdesen (1992)	1967:3– 1989:4 Quarterly	J (1988); JJ (1990)	M	1.374	0.810	6.553	-1.544 & -0.995	-0.0978
Switzerland Chowdhury (1995)	1973:2– 1991:4 Quarterly	J (1988); JJ (1990)	b mI	0.940 0.887			-0.260 -0.310	0.363 & -0.140 ⁹ 0.344 & -0.098 ⁸ 0.391 & -0.102 ¹⁰ 0.308 & -0.052 ¹⁰
United Kingdom Drake and Chrystal (1994)	1976:2– 1990:3 Quarterly	J (1988); JJ (1990)	Mld Mld M2d M3d	3.223 3.372 2.560 2.576	1.041 0.815 1.208 1.190	0.775 -0.707 1.087 -0.769	-0.032 -3.765 -4.187	-4.346 ¹¹ -4.829 ¹¹
United States Miller (1991)	1959:1– 1987:4 Quarterly	EG (1987)	M2	1.204	0.952		-0.092*	
Baba, Hendry, and Starr (1992)	1960:3– 1988:3 Quarterly	J (1988)	mI	0.510			-6.640	-5.510 & 3.720 ¹²
McNown and Wallace (1992)	1973:2– 1988:4 Quarterly	J (1988); JJ (1990)	mI mI m2 m2	0.987 1.001 1.131 1.128	0.987 1.001 1.131 1.128		-2.828 -9.600 -1.745 -1.747	0.133* ⁷ 0.131* ⁷

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Table 2. (continued)

Study	Period/ Frequency	Method	Money ³	Elasticity		Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴
Developing countries								
Argentina Choudhry (1995)	1935:1– 1962:4 Quarterly	J (1988); JJ (1990)	<i>m</i> <i>m</i> <i>m</i> <i>m</i>	1.970 1.680 1.910 3.450			-0.025 -0.033 -0.034 -0.041	
Cameroon Fielding (1994)	1977:1– 1987:2 Quarterly	JJ (1990)	<i>m</i> <i>m</i>	1.490			-8.910* -1.310*	-8.100 ¹³
China Tseng and others (1994)	1983:1– 1988:4 Quarterly	EG (1987)	<i>c</i> <i>c</i> <i>m</i> <i>m</i>	1.900 1.530 1.810			-1.230 -1.510 -2.210	
Côte d'Ivoire Fielding (1994)	1975:3– 1987:4 Quarterly	JJ (1990)	<i>bm</i>	1.580			-0.030 -0.940 -1.540 -0.050	2.430* -3.040* -1.630 ¹³

Table 2. (continued)

Study	Period/ Frequency	Method	Money ³	Elasticity			Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴	Inflation
									Other
India Moosa (1992)	1972:1– 1990:4 Quarterly	EG (1987)	cc mm bm	0.874 0.785 1.471				-0.109*	
		J (1988) JJ (1990)	cc mm bm	0.986 0.797 1.573				-0.032* -0.172*	
Indonesia Price and Insukindro (1994)	1969:1– 1987:4 Quarterly	EG (1987)	chp dd	0.880 1.300				-2.100 ⁸ -1.000 ⁸	
		J (1988); JJ (1990)	chp dd	0.710 1.100				-4.400 -8.400	
Iran Bahrani– Oskooee (1996)	1959–90 Annual	J (1988); JJ (1990)	m2 m2	1.390 1.330				-3.300 ⁸ -9.100 ⁸	
Kenya Adam (1992)	1973:1– 1989:2 Quarterly	J (1988); JJ (1990)	m0 m1 m2 m3 m3d	1.010 0.890 0.840 1.100 0.840				-6.150 -5.460 -6.730 -6.190 -5.510	
								-0.160 ¹⁴ -0.110 ¹⁴ -0.090 ¹⁴ -0.070 ¹⁴	

Table 2. (continued)

Study	Period/ Frequency	Method	Money ³	Elasticity		Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴
Korea Arize (1994)	1973:1– 1990:1 Quarterly	EX(1987)	<i>mI</i>	0.500			-0.027	-0.007 & -0.016 ¹⁵
			<i>m2</i>	0.950			-1.220	-0.003 & -0.080 ¹⁵
		J (1988); JJ (1990)	<i>mI</i>	0.570			-0.034	-0.008 & -0.021 ¹⁵
			<i>m2</i>	1.160			-9.150	-0.017 & -0.090 ¹⁵
Lebanon Eken and others (1995)	1964–93 Annual	EG (1987) PO (1990)	<i>b</i>	0.790			-1.200	
			<i>mI</i>	1.120			-1.470	
			<i>m2II</i>	0.960			-1.310	
Malaysia Sri Ram (1999a)	1973:8– 1995:12 Monthly	J (1988); JJ (1990)	<i>m2</i>	1.036			4.884 -5.391	-4.745 -4.891 -0.581 ⁷
			<i>m2</i>	1.130			2.510 -1.834	
Mexico Khamis and Leone (1999)	1983:1– 1997:6 Monthly	J (1988); JJ (1990)	<i>cc</i>	0.450			-9.730	

Table 2. (continued)

Study	Period/ Frequency	Method	Money ³	Elasticity			Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴	Inflation
Morocco Hoffman and Tahiri (1994)	1959:1– 1988:2	OLS	<i>M1</i>	1.080 1.120	1.330 1.290		-0.050 -0.050		-0.020* -0.030*
	Quarterly		<i>M2</i>	1.100 1.140	1.120 1.090				-0.010* -0.020*
			<i>DOLS</i>	<i>M1</i> 1.200	1.080 1.050		-0.025 -0.020		-0.040* -0.040*
			<i>M2</i>	1.210 1.230	1.900 0.900				-0.030* -0.040*
			<i>J</i> (1991b)	<i>M1</i>	1.120 1.100	0.940 0.970			-0.060* -0.060*
			<i>M2</i>	1.180 1.170	0.860 0.870				-0.040* -0.040*
			<i>JJ</i> (1990)	<i>bm</i>	0.720				-4.430 ¹³
	Nigeria Fielding (1994)	Quarterly				1.180*	-1.420*		
Tertiba (1997)	1960–94	EG (1987)	<i>COP</i> <i>M1</i> <i>M2</i>	1.325 1.525 1.317	1.057 1.051 0.626	2.683* 2.854* 2.859*	-2.854* -2.819* 2.122*	-2.854* -2.819* -2.209*	-0.314* ¹⁶
	1962:1– 1995:2	Quarterly	<i>EG</i> (1987)	<i>M1</i>	1.607	0.843	0.663*		-0.286* ¹⁶
			<i>EG</i> (1987)	<i>M2</i>	1.146	0.269	0.943*	-0.592*	

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Table 2. (continued)

Study	Period/ Frequency	Method	Money ³	Elasticity		Opportunity Cost (Semi-Elasticity) ²		
				Real Income	Price Level	Own-Rate	Alternative Return	Interest Rate ⁴
Pakistan Arize (1994)	1973:1– 1990:1 Quarterly	EX(1987)	<i>m</i> <i>m</i> <i>m</i> <i>m</i> <i>m</i> <i>m</i> <i>M</i>	0.930 0.990 1.030 0.770 0.860 1.070 N/M			0.003 0.038 -0.540* -0.050*	-1.130 -1.270 -5.480 -7.880
Hossain (1994)	1972–91 Annual	J (1988); JJ (1990)	<i>m</i> <i>m</i>					-0.030 ¹⁴ -0.023 ¹⁴
Singapore Arize (1994)	1973:1– 1990:1 Quarterly	EX(1987)	<i>m</i> <i>m</i> <i>m</i> <i>m</i> <i>m</i> <i>m</i>	0.720 1.080 0.710 1.120 0.620 1.620			-0.330 -0.030 -0.110 -0.030	-1.790 ¹⁷ -1.980 ¹⁷
Dekle and Pradhan (1997)	1975–95 Annual	J (1988); JJ (1990)						-1.780 ¹⁷ -1.830 ¹⁷
Thailand Dekle and Pradhan (1997)	1978–95 Annual	J (1988); JJ (1990)	<i>N</i> / <i>M</i>	1.130		0.670		-0.009
Tunisia Treichel (1997)	1990:1– 1995:12 Monthly	J (1988); JJ (1990); AEG	<i>m</i> <i>m</i> <i>m</i>	0.130 1.070 0.800			-0.020 -0.030 -0.008	

Table 2. (concluded)

- ¹Refer to Table 1 for corresponding expansion on abbreviations used in this table.
- ²Semi-elasticities except for those market by * which refer to elasticities.
- ³Variables in nominal term are shown in upper case letters and in real term in lower case; and all variables are in italics to show that they are expressed in logarithmic term.
- ⁴Where own-rate or alternative return is not explicitly stated; also refers to the net interest rate measure.
- ⁵Financial innovation variable.
- ⁶Elasticities of those variables expressing the income and wealth concepts respectively.
- ⁷Exchange rate measure.
- ⁸A measure of foreign interest rate.
- ⁹Short-term interest rate for alternative return, and the other category includes both NEER and a measure of foreign interest rate.
- ¹⁰Long-term interest rate for alternative return, and the other category includes both NEER and a measure of foreign interest rate.
- ¹¹Implicit divisia rental price or user cost index.
- ¹²Financial innovation variable and volatility measure for yield on long bond.
- ¹³Measure of price variability.
- ¹⁴Exchange rate depreciation.
- ¹⁵Foreign exchange risk and a measure of foreign interest rate respectively.
- ¹⁶Spread between local and foreign interest rates.
- ¹⁷Foreign exchange risk.
- ¹⁸Variable expressing foreign influence.

Figure 1. Frequency Distribution of Estimated Income Elasticities for Components of Narrow Money

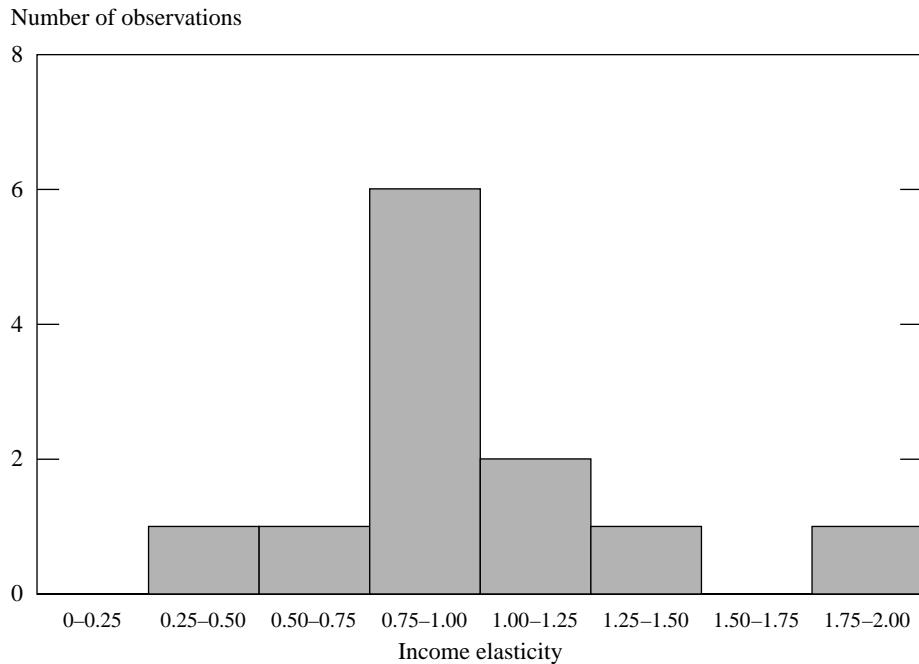


Figure 2. Frequency Distribution of Estimated Income Elasticities for Components for Narrow Money

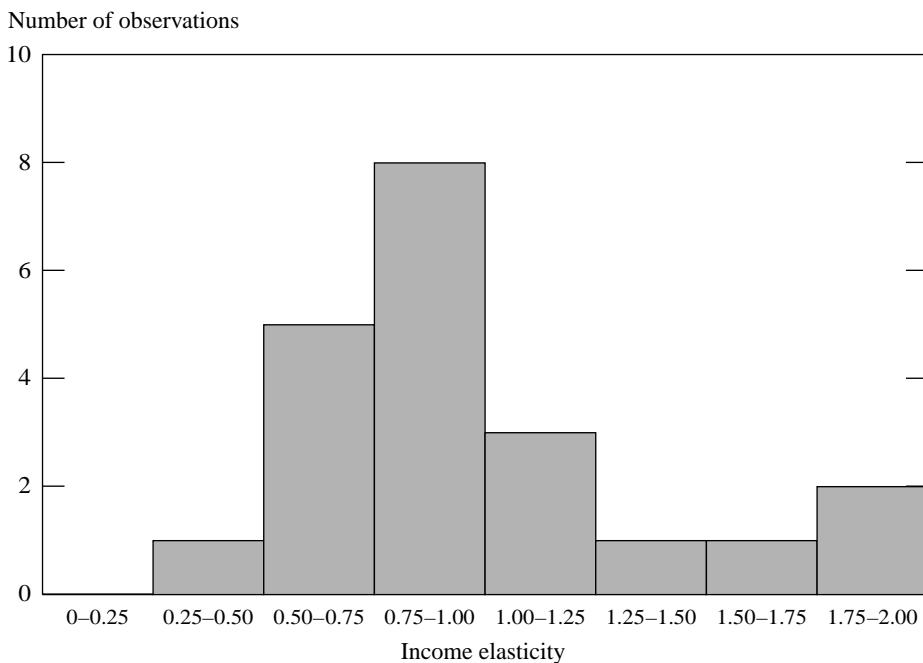


Figure 3. Frequency Distribution of Estimated Income Elasticities for Components for Broad Money

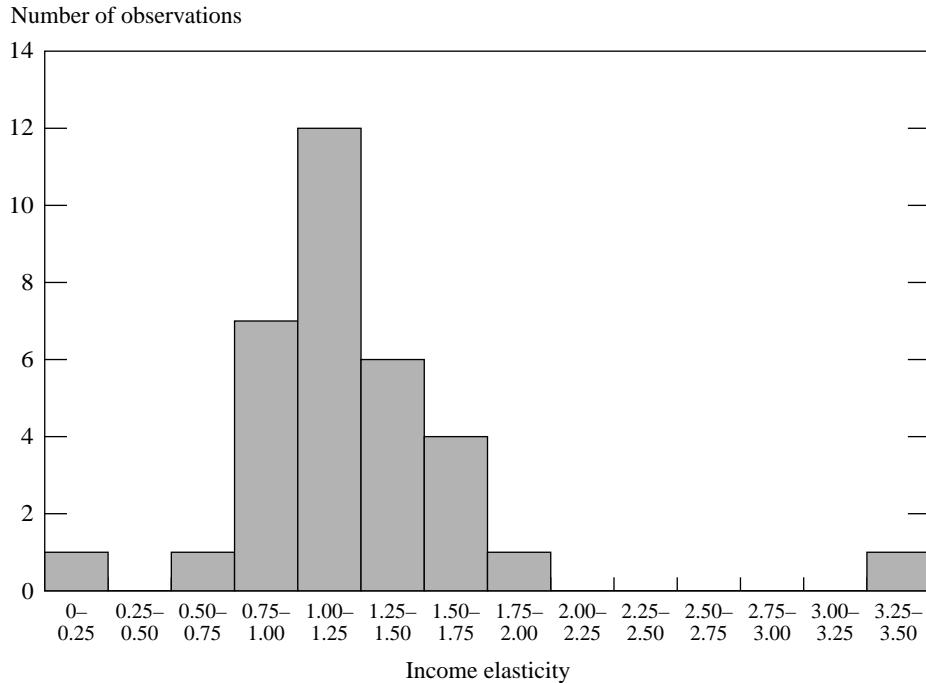


Table 3. Descriptive Statistics for Income Elasticities

	Number of Observations	Mean	Median
Components of narrow money	12	0.99	0.95
Narrow money	21	0.98	0.89
Broad money	33	1.22	1.13

Source: Table 2.

inflation. As can be seen from Tables 1 and 2, there are a number of other variables considered to tackle the country-specific issues; in addition, the open-economy type models also employ the foreign opportunity cost variables.

III. Conclusion

The study has made an attempt to survey a number of papers that applied the error-correction models to analyze the demand for money in a number of industrial and developing countries. The major contribution of this paper is that it has summarized the major features of these papers and presents relevant information in a comparable framework to promote easy understanding of the approaches followed, variables included, and coefficients derived. The information presented

thus will enable the researchers to compare their own results and approaches with what were undertaken previously in a wide range of countries. Alternatively, it will help identify important factors to be considered before modeling and estimating money demand in other countries exhibiting similar or different economic characteristics. In short, it will provide a starting point to conduct the money demand research using the error-correction approach.

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