# Market Power in the Middle East

Yevgeniya Korniyenko, Ahmed Tohamy, and Weining Xin

WP/25/1

IMF Working Papers describe research in progress by the author(s) and are published to elicit comments and to encourage debate.

The views expressed in IMF Working Papers are those of the author(s) and do not necessarily represent the views of the IMF, its Executive Board, or IMF management.

**2025** JAN



#### **IMF Working Paper**

Middle East and Central Asia Department

# Market Power in the Middle East Prepared by Yevgeniya Korniyenko, Ahmed Tohamy, and Weining Xin\*

Authorized for distribution by Amine Mati January 2025

*IMF Working Papers* describe research in progress by the author(s) and are published to elicit comments and to encourage debate. The views expressed in IMF Working Papers are those of the author(s) and do not necessarily represent the views of the IMF, its Executive Board, or IMF management.

ABSTRACT: The Middle East (ME) is often perceived as region with rentier economies and uncompetitive markets. Evidence of market power in the region however is scant. In this paper, we ask the following three questions: Is the ME uniquely uncompetitive? Has the evolution of market power in the region traced the global rise in market power? What government policies and actions influenced the market power in the region and can taxes be a way to even the playing field? To answer these questions, we utilize comprehensive firm-level data from Compustat between 2004 and 2022 and employ two methods for estimating markups (production function and cost-share approach). We document that market power among listed firms in the ME is higher than in the US, but on a downward trend. We find that the value-added tax (VAT) reforms introduced by some Gulf states from 2018 to 2022 resulted in a reduction of market power, an additional benefit beyond increasing fiscal space. While policymakers should continue to use available regulatory levers to achieve economic efficiency and a level playing field, VAT could be considered as an alternative instrument.

| JEL Classification Numbers: | D4, H2, L1, O3, N4   |
|-----------------------------|--|
| Keywords:                   | Middle East; Market Power; Markup; Firm Behavior                     |
| Author's E-Mail Address:    | ykorniyenko@imf.org, ahmed.tohamy@nuffield.ox.ac.uk,<br>WXin@imf.org |

<sup>\*</sup> The author(s) would like to thank Johannes Abeler, Steve Bond, Amine Mati, Jerome Vacher, Sidra Rehman, Usama Jamal and Giulio Gottardo for helpful discussions. We also thank participants in the divisional presentation in the Middle East & Central Asia Department (MCD) of the IMF and in the MCD Hybrid Discussion Forum for helpful comments. The views expressed in the paper are those of the authors and do not necessarily represent the views of the IMF, its Executive Board, or IMF management.

#### **WORKING PAPERS**

# **Market Power in the Middle East**

Prepared by Yevgeniya Korniyenko, Ahmed Tohamy, and Weining Xin

# **Contents**

| I. I  | Introduction  | 6  |
|-------|---|----|
| II.   | Background Literature   | 8  |
| III.  | Methodology   | 9  |
| IV.   | Firm-Level and Macroeconomic Data   | 11 |
| V.    | Main Results  | 12 |
| VI.   | Conclusion  | 22 |
| Refe  | erences   | 24 |
| Ann   | ex I. Sample Representativeness   | 27 |
|       | A.I.a. Earnings relative to GDP   | 27 |
|       | A.I.b. Sales relative to GDP  |    |
| Ann   | ex II. Sample Description   | 29 |
|       | A.II.a. Table 1: Summary Statistics for Sample between 2004 and 2022              | 29 |
|       | A.II.b. Sample Composition: by Firm Location                                      | 29 |
|       | A.II.c. Sample Composition: by Country of Incorporation                           | 30 |
|       | A.II.d. Sample Composition: by Industry   |    |
| 4     | A.II.e. PFA: excluding XSGA   | 31 |
| Ann   | ex III. Revenue-Variable Costs Ratio (CSA)  | 32 |
| Ann   | ex IV. Value Added Tax: Robustness Table  | 33 |
|       | ex V. Country Specific Results  |    |
|       | A.VI.a. Saudi Arabia  |    |
|       | A.VI.b. Egypt   | 35 |
| Ann   | ex VI. Breakdown of Markups Dynamics  | 36 |
| Ann   | ex VII. Correlation between Markup Measures and HH-index Measure of Concentration | 37 |
| Ann   | ex VIII. Additional Results   | 38 |
| FIGL  | JRES  |    |
| Figui | re 1. The ME Markup: Different Weights  | 12 |
| Figui | re 2. Sales-Weighted Average ME and US Markups                                    | 13 |
| Figu  | re 3. Markups Estimates Using Different Methods (excluding ARAMCO)                | 13 |
| _     | re 4. The ME Superstar Firm Phenomena   |    |
| 3,50  | 1   |    |

| Figure 5. Map of Markups in the ME in 2010-2019  | 14 |
|--|----|
| Figure 6. Markup Heterogeneity by Country  | 14 |
| Figure 7. Markups Heterogeneity by Industry  |    |
| Figure 8. Markups for domestic vs external oriented firms in ME                              | 16 |
| Figure 9. ME Profits and Dividend Payouts  | 17 |
| Figure 10. Markup Impulse Response Function (IRF) to an Inflation Shock                      | 18 |
| Figure 11. VAT Adoption by Gulf states   | 19 |
| Figure 12. Markups in GCC Countries That Adopted VAT vs Countries That Did Not               | 20 |
| Figure 13. Markup Wedge Between Final Goods and Intermediate Goods upon the Tax Introduction | 21 |
| Figure 14. Staggered Diff-in-Diff Estimator for Impact of VAT Adoption on Markups in the GCC | 22 |

#### I. Introduction

The Middle East (ME) is often perceived as a region with rentier economies and uncompetitive markets. The perception is attributed to several factors, including: a large dependence on natural resources and commodities revenue (in particular in Gulf Cooperation Council (GCC) countries) despite decades of policy efforts to diversify the economies; the presence of significant market distortions; and, a large presence of State-Owned Enterprises (SOEs) that have weakened efforts to develop a competitive private sector; and a need for stronger governance and institutions (Chaudhry, 1997).<sup>1</sup>

In this paper, we ask a few main questions: Is the Middle East uniquely uncompetitive? How has market power evolved in the region in light of multiple reforms? Are GCC countries' markups different from the rest of the Middle Eastern countries? Has the region seen a marked trend of rising market power similar to the rest of the world? Did markups add to inflationary pressures in GCC and the Middle East? Additionally, we answer the question of whether, in a context with limited capacity to enforce anti-trust laws, value-added tax (VAT) policy can act as a backstop to the role of antitrust with respect to inducing competition. Answering this question has serious implications for the role and importance of antitrust, taxation, and product market reforms in the region. While research on market power has gained prominence recently, few studies have examined market power and its evolution in the Middle East thus far (Robinson and Acemoglu, 2012).<sup>2</sup> In contrast, a growing body of research has identified a trend of increasing corporate market power both in the US and across the world (De Loecker et al., 2020, Diez et al., 2018, Autor et al., 2020).

To answer these questions, we employ the production function approach (pioneered by De Loecker & Warzynski (2012)) and the cost-share approach (as suggested by Bond and Gottardo, (Forthcoming)) to estimate firm-level corporate markups for the Middle East between 2004 and 2022 using comprehensive publicly listed firm accounts from Compustat.<sup>3,4</sup>

We find that corporate market power among listed firms in the Middle East is indeed higher than in the US, but it is on a downward trend (except for the COVID-19 period).<sup>5</sup> Like the rest of the world, the Middle East is experiencing a superstar firm phenomenon, which explains two simultaneous trends: the rise in market concentration and the fall in labor shares. Within the region, GCC countries have higher corporate markups (including or excluding oil companies) than the rest of the Middle East countries in our sample, and their

<sup>&</sup>lt;sup>1</sup> This perception has frequently fed into policy decisions relating to the region. The most recent example is the IMF's reform recommendation on Egypt to strengthen its antitrust laws and to release a government strategy for an `SOE-exit' from economic sectors.

<sup>&</sup>lt;sup>2</sup> Robinson & Acemoglu (2012) provide a historical account of the development of the market economy in the region during the Ottoman empire. Additionally, Diwan et al. (2020) highlighted the prevalence of crony activities by politically connected firms in the Middle East, which hindered competition and impeded economic growth. In fact, Acemoglu et al. (2017) suggest that the Arab spring protests served as a partial check on rent-seeking and wonder if the protests were caused by rent-seeking. None of these estimate market power in the region and contrast it to estimates in the US or the EU.

<sup>&</sup>lt;sup>3</sup> Markups are defined as the ratio of firms' prices to their marginal cost. In a neoclassical world with perfect competition firms price their products equal to marginal cost. Markups then measure the extent to which firms deviate from this competitive environment.

<sup>&</sup>lt;sup>4</sup> We use the cost-share approach to avoid the critiques documented in the literature by Bond et al. (2021). To ensure the robustness and validity of the results for Middle Eastern countries, we expand our sample to non-listed companies using the BvD Orbis data and re-run the estimates using both estimators. We also provide alternative measures of market power including market concentration, for example.

<sup>&</sup>lt;sup>5</sup> When comparing our estimates with the literature on European firms, we find that the aggregate markups we have in the Middle East are higher than those reported by Loecker & Scott (2017). Loecker & Scott (2017) is the closest method to our implementation.

markups shot up immediately after the pandemic but are converging to the US and the rest of the Middle East since then. Our main results are robust to the choice of markup estimator. The cost-share markup estimator requires the additional assumption of constant returns to scale. However, it relaxes the need to identify the output elasticity of production which is heavily disputed in the literature.

Against the backdrop of global inflation during 2021-22, we examine whether inflation has been passed on to consumers. Our findings suggest that GCC countries appear to have absorbed some of the price rises by lowering their markups following inflation surprises, unlike other countries in the region. The significance of this result varies by country, but the relationship remains negative. Generally, we do not find evidence of "greedflation" in the Middle East.

Since our sample coverage expands over the period analyzed, it is possible that the (net) entry of more productive firms drives the decline in our markup measures. To rule out this possibility, we demonstrate that within-firm markups are falling and driving a significant portion of the decrease in aggregated weighted markups for both measures. According to one of the markup measures, net entry also contributes to the falling markups over the period. However, for the other measure, net entry plays a very limited role in the evolution of markups in the region.

We also examine whether there is any heterogeneity across sectors and across firms with different degrees of international market exposure. Companies operating in marketable services sectors and oil and utilities industries in the Middle East typically have higher markups relative to other industries. While firms with zero foreign currency exposure (more domestically focused) generally have lower markups, the overall dynamic of markups is driven by firms with larger foreign exposures (particularly for non-energy companies). This has two implications. First, lower corporate markups of domestically oriented firms are benefiting domestic consumers in the Middle East. Second, the reduction in markups is likely attributable to firms in the region gaining global competitive efficiency, which could have positive productivity externalities on the rest of the ME firms.

In terms of policies that influenced market power, our findings suggest that first-generation trade and financial market reforms, as well as improvements in property rights in ME region, have contributed to the falling markups trend. To examine whether VAT policy can serve as a backstop to antitrust efforts in promoting competition, we utilized the cross-country variation in VAT adoption in GCC. Our results indicate that the VAT reforms introduced and adjusted by some GCC countries from 2018 to 2022 led to a reduction in market power, an additional benefit beyond increasing fiscal space.

The rest of the paper is organized as follows. Section 2 provides a background literature overview. Section 3 describes the markup estimation methodology and the challenges in adopting the estimation procedure, including the costs and benefits of adopting an alternative method described in Annex III. Section 4 describes the sample and data source. Section 5 discusses the results, including stylized facts about market power in the GCC and the rest of the Middle East and their macroeconomic implications. Section 6 concludes.

## II. Background Literature

This paper relates and contributes to several strands of literature. First, this study contributes to the literature on rent-seeking in the ME region by providing comprehensive firm-level estimates of market power in the region. The estimates of firms' markups and their dynamics can help determine whether the Middle East truly exhibits a unique lack of competition, which, in turn, can inform the necessity for strengthening antitrust laws and regulations to ensure a level playing field for all market participants (Gigineishvili and others, 2023, IMF Middle East and Central Asia (ME&CA) REO, 2023). This study enhances the understanding of the region's market competition and competitiveness.

Our research also makes a significant contribution to the literature on the development of global corporate market power, as evidenced by the works of Akcigit et al. (2021), Qureshi (2019), and Diez et al. (2018). Specifically, this study aims to assess whether a significant regional economic bloc, such as ME, is experiencing a similar trajectory of rising corporate market power as the rest of the world. Earlier qualitative work suggested such a trend for the region (Diwan et al. 2020). Building upon this, our quantitative analysis adds empirical evidence to the discussion, while also challenging previous views.

Third, to corroborate our estimates and address criticisms related to markup estimation using the production function approach (De Loecker et al., 2017), we employ an alternative adjusted markup estimator. This methodology addresses concerns raised in recent work by Bond et al. (2021), Raval (2023), ensuring the robustness and reliability of our findings.

Our research also makes a significant contribution to the literature on the macroeconomic implications of corporate market power in the ME and the GCC. The recent rise in global inflation is often attributed to corporate greed, which is seen as a key driving force behind rising prices. Therefore, from a policy perspective, understanding the contributing role of the dynamics of market power to inflation is important. We use the markup estimates for ME countries and relate them to headline inflation in the region. Our analysis suggests that a rise in corporate market power did not contribute to the inflation surge in GCC countries in 2022. These findings are consistent with the findings for advanced economies (Chapter 1, October 2023 World Economic Outlook).

Besides documenting the dynamics of market power in the ME, we also explore policies that could enhance competition and contribute to the economic diversification of the region, adding to the literature on economic diversification and competition. As far as we know, we are the first to study the market power implications of value-added taxes. We argue that tax policy, particularly VAT, could be used to deter market power in the region and promote competition. Our analysis suggests that VAT reforms have had additional benefits, such as reducing market power, in addition to increasing fiscal space in GCC countries. While policymakers should continue to use available antitrust levers to deter market power and achieve economic efficiency, VAT could be a novel additional instrument to consider. This builds up on theoretical work by Delipalla & Keen (1992) on the efficiency of ad valorem taxes. It also builds on empirical work by Benzarti et al. (2020).

# III. Methodology

#### III.a. Markup Estimation: Production Function Approach (PFA)

A sizeable literature starting from Hall (1988) recognized that price-marginal cost markups can be recovered as

$$\mu_{it} =: \frac{P_{it}}{MC_{it}} = \epsilon_{it} \cdot \left[ \frac{S_{it}}{R_{it}} \right]^{-1},\tag{1}$$

where  $\epsilon_{it}$  is the output elasticity of the variable input s,  $S_{it}$  is the expenditure on input s, and  $R_{it}$  is sales or turnover. This follows from the cost minimization problem of a firm with variable costs described by :

$$\min_{s} C(s) \quad s.t. \ Q = F(s) \quad \& \quad Q \ge 0$$

where C(s) is the cost function of a vector of inputs, Q is the quantity of output produced, and F(s) is the production function of Q. The first order condition is thus:

$$P_{S} = \lambda \frac{\vartheta Q}{\vartheta S}$$

where Ps is a vector of containing the price of the inputs s,  $\lambda$  is the shadow cost of an extra unit of output (the marginal cost), and  $\frac{\vartheta Q}{\vartheta s}$  is a vector of the derivative of quantity produced with respect to the variable inputs s. Dividing and multiplying by P (the price of the output good) on the right hand side and rewriting  $\frac{\vartheta Q}{\vartheta s}$  as an elasticity, one obtains equation 1.

Equation 1 has attractive features for estimating markups. Firstly, we do not require any data relating to the demand schedule of the output product market. Secondly, we only require data on expenditure share on a variable input, s, relative to total sales and an estimate for the output elasticity of the same input. There are challenges however in estimating  $\epsilon_{it}$ . We typically observe revenue data and, hence, can only estimate revenue elasticity,  $\eta_{it}$ , as opposed to output elasticity,  $\epsilon_{it}$ . For example, De Loecker & Warzynski (2012), De Loecker et al. (2020) use revenue measures from Compustat to estimate corporate markups, based on equation 1. This departure from using output elasticities leads to challenges to the identification of the markup level (Bond et al., 2021). The literature usually places restrictions on the relationship between  $\eta_{it}$  and  $\epsilon_{it}$ . These concerns are discussed in the next subsection.

Note that the production function approach is in contrast to the structural approach in the literature that estimates a demand schedule and uses that to back out industry- or firm- level markups (Syverson, 2019, Berry et al., 1995).

To estimate  $\eta_{it}$ , in our baseline specifications, we adopt the following auxiliary regression for  $\eta_{st}$ .

$$y_{it} = \eta_{st}^{v} v_{ist} + \eta_{st}^{x} x_{ist} + \eta_{st}^{k} k_{ist} + \omega_{ist} + u_{ist}, \tag{2}$$

where lowercase letters denote logs,  $y_{it}$  is logged realised firm revenue assuming  $y_{it} = ln(Q_{it}exp(u_{it}))$ ,  $v_{ist}$  is the log of a composite of cost of goods sold,  $x_{it}$  is the log of a composite measure of overhead expenses,  $k_{it}$  is the log of firm's capital stock,  $\omega_{it}$  is the log of a hicks neutral productivity shock. Relying on the insight of Olley & Pakes (1996) that productivity,  $\omega_{it}$ , can be expressed as a semi-parametric function of a firm's state variables and observables, we model  $\omega_{it}$  as

$$\omega_{it} = \phi_{st}(i_{ist}, k_{ist}, z_{ist}) + u_{ist}, \tag{3}$$

These restrictions allow us to estimate equation 2 as a simple semi-parametric regression. Assuming  $v_{ist}$  is a variable cost, we can then use  $\eta^v_{st}$  to recover markups. To identify  $\eta^v$ , we need to assume that the variable cost bundle, v, is non-dynamic and chosen at t, after the investment decision, made at t-1, while allowing productivity shocks to hit the firm in between those two periods. Aside from its easy implementation, there are two reasons we adopt this approach. Firstly, this (control function) approach is not very data-demanding relative to GMM methods like Ackerberg et al. (2015). This is important given our small sample as discussed later. Secondly, in practice as confirmed in previous analysis for the US (De Loecker 2021), the difference in revenue elasticities estimated based on competing approaches such as Ackerberg et al. (2015) is not large.

In our baseline specification, we estimate a unique,  $\eta_{st}$  for each sector, holding it constant across the entirety of the sample again in order to increase our effective sample size. This effectively assumes that the production function is constant over the sample period,  $\eta_{st}=\eta_s$ . We also run a variety of estimation techniques to confirm our results. Those are discussed in the Annex II.

### III.b. Challenges to the Adopted Estimation Procedure

- Sample Size. Our research aims to estimate the output elasticity  $\eta_s$  of firms in the region, which has few listed firms. To reduce the data requirements for the estimation procedure, we adopt multiple solutions. Firstly, we estimate the output elasticity assuming  $\eta_s$  varies across more aggregated sectoral classifications, instead of the finer 2-digit sectoral classifications used in the literature. This increases the statistical power of the estimator in the auxiliary regression. Secondly, we assume the production function is constant over the sample period to allow the sample to be larger. We use an estimation procedure that relies on less data restrictions. Lastly, we expand the sample size by including non-listed firms from BvD Orbis database to estimate the same parameters for non-listed firms. This is one of the few papers that has expanded in this direction, the other being Raval (2023). While we achieve a larger sample, we also introduce more heterogeneity by expanding the sample to include unlisted firms, which needs to be treated with caution<sup>6</sup>.
- **Methodology.** There are two further critiques levelled against the production function approach. Firstly, Raval (2023) finds that production function estimation based on different inputs yields different markup distributions. This fact may be related to the necessary assumption that the procedure ought to use only

<sup>&</sup>lt;sup>6</sup> To study the relevance of our results for an unlisted sample of firms across both the ME and the US we ran a PFA markup estimator using ORBIS data. Note XSGA is not reported in ORBIS & so we cannot compute PFA in our baseline. We construct PFA (excluding XSGA) similar to the previous Annex II. Note we also cannot construct the share estimator given the absence of XSGA.

variable costs. In practice, some costs used to estimate markups might not be variable, leading to this discrepancy. Secondly, firm accounts do not feature output data but rather revenue data. Hence, the estimated markups are based on revenue elasticities. This leads to a downward bias in the estimated markups as suggested by Klette & Griliches (1996). Additionally, Bond et al. (2021) outline a host of problems in estimating markups when only revenue and expenditure data is available. These criticisms are generally because of the omitted price bias concerns. In Annex II.e., we study the impact of estimating of placing different assumptions on the estimation of the revenue elasticity to explore robustness of the main results reported.

#### III.c. Markup Estimation: Cost-Share Approach (CSA)

We additionally detail an alternative cost-share approach to estimating markups that relies on the key assumption of constant returns to scale. The benefit of this approach is that it relaxes the need to identify the widely disputed output elasticity using the production function approach relying on revenue accounting as opposed to production data.

Assuming cost minimization and constant returns to scale, we can show that an estimate for the output elasticity for input s,  $\epsilon^{s}_{it}$ , is

$$\epsilon_{it}^{s} = \frac{\mathrm{E}[\omega_{s} \cdot s_{it}]}{\mathrm{E}[\Sigma_{i} \omega_{i} \cdot j_{it}]}.$$

Based on this result we can recover markups using equation 1 as:

$$\mu_{it} = \frac{R_{it}}{\mathbb{E}[\Sigma_j \, \omega_j \cdot j_{it}]}.\tag{4}$$

Equation 4 intuitively suggests that markups are identified as the ratio of revenue to the expected value of variable costs. This alternative method places an additional assumption of constant returns to scale. However, the benefit is this method relaxes the need to identify output elasticity using the production function approach relying on expenditure data. Bond et al. (2021) emphasize the need for quantity data to identify output elasticity for markups and also highlight that using expenditure data instead leads to biased markups. Annex III presents detailed results using CSA method.

### IV. Firm-Level and Macroeconomic Data

We use comprehensive consolidated accounts of publicly listed firms from Compustat Global restricting our sample to firms whose main headquarters is in one of 13 ME economies (including all six Gulf states) between 2000 and 2022<sup>7</sup>. Annex I reports the coverage of the sample relative to the size of these economies over time and by country using profits and sales data.

We deflate firm accounts using country-level deflators from the IMF World Economic Outlook (WEO) database and convert them to USD in order to allow for cross-country comparison. In line with the literature (for e.g.

<sup>&</sup>lt;sup>7</sup> The sample also includes Kazakhstan.

Hennessy & Whited (2005), we drop observations with negative sales or cost of goods sold. We also winsorize the ratio of sales-to-costs of a goods sold at the 1 percent double-sided in order to remove outliers. This results in a sample of approximately 1300 firms and 20,700 firm-year observations. Annex II reports summary statistics on sales, cogs, profits, firm-year observations by country; the country composition of the firms' location and the industry composition of these firms per country.

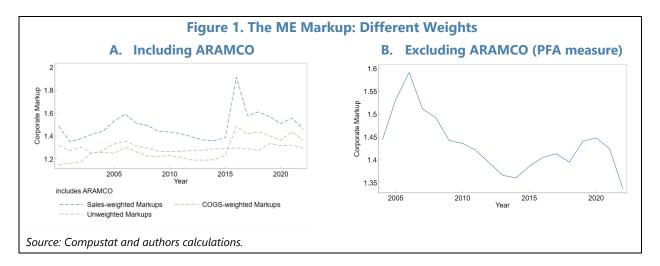
To benchmark the results, we construct a sample of US-incorporated publicly listed firms from Compustat in order to compare markups in the region with those for the US firms. We apply the same sample restrictions for this sample of US firms.

Other macroeconomic variables, such as gross domestic product, production by industry, inflation, and value-added tax (VAT) rates, are utilized to construct our variables. These are available from the IMF WEO database.

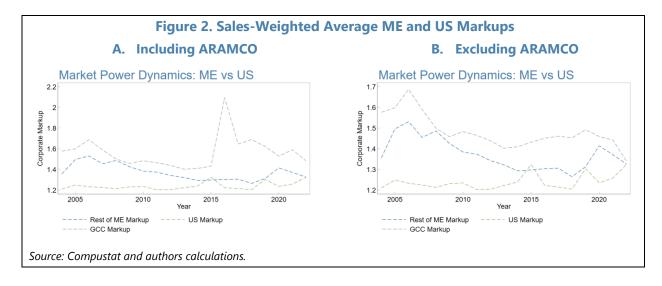
### V. Main Results

#### V.a. Market Power in the Middle East

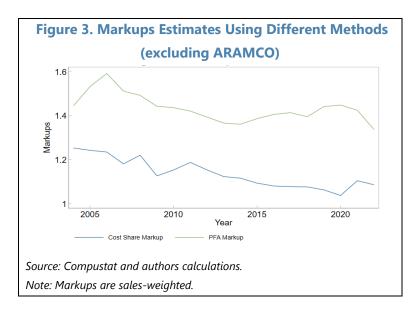
Market power in the GCC and the rest of ME region has been higher than in the US.8 This result is consistent with other measures of markups, profit shares, dividend payouts, and market concentration (see Annex II-IV, VII, VIII). However, unlike the US, GCC market power has not shown an upward trend over the last twenty years. A significant spike in market power in 2016 is observed after the listing of Saudi Arabia's ARAMCO. Despite the rising global trends of corporate market power, the GCC and the ME have escaped the trend (Figure 1 and 2).



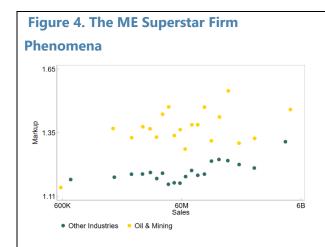
Our markup estimates for the ME&CA are also higher than those reported in the literature for European countries (Loecker & Scott 2017).



In a perfectly competitive market, markups should be equal to 1. Deviations from 1 indicate market power. The average markup in the GCC was close to 1.6 around 2010. Excluding ARAMCO, the trend is downward and stable around 1.45 between 2010 and 2020. To mitigate the criticism summarized by Bond et al. (2021) on the bias of markup estimation in the absence of quantity data, we estimate markups for the ME using an alternative method (Figure 3). This method, described in Annex A.II.e., assumes constant returns to scale, which imposes additional structure on the problem and obviates the need for quantity data. The alternative method reveals a similar dynamic for the ME region, with markups decreasing over time.

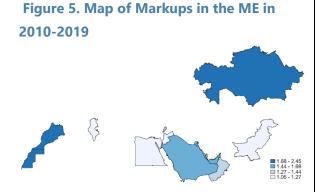


Listed firms in the ME display a "superstar" phenomenon, where firms with higher sales have higher market power. This is true for both the oil, mining, and utilities sector and other sectors (Figure 4). The oil, mining, and utilities sector has higher market power than other sectors across the entire sales distribution, indicating some unique market dynamics associated with OPEC+. Resource-rich countries in the ME are also the ones with higher average markups. Morocco is the only non-oil country in the sample with comparable level of markups (Figure 5).



Source: Compustat and authors calculations.

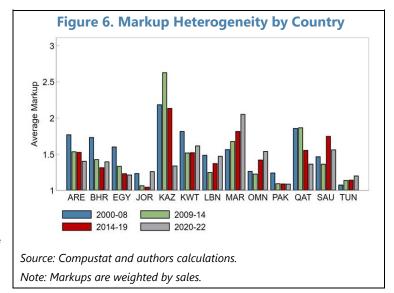
Note: Companies in Oil, Mining, & Utilities sector are shown as yellow dots and companies from all other Industries are represented by teal dots. Logged scale.



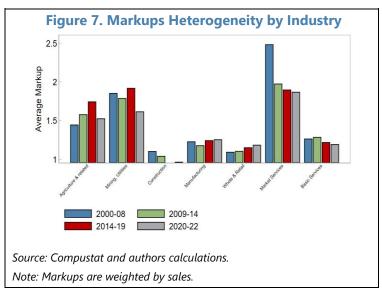
Source: Compustat and authors calculations.

Note: Sales-weighted Markups by country in the Middle
East. Darker blue color showing higher market power.

Within the GCC, average markups did not change much over the period, but they showed a weak upward trend in Bahrain, Kuwait, and Oman in recent years (Figure 6). Conversely, Qatar, Saudi Arabia, and the UAE showed a downward trend. This fact can have multiple explanations. First, one can argue that some countries succeeded in implementing structural reforms to enhance market competitiveness and strengthen antitrust laws. The passage of several laws across the region, such as Qatar's competition law of 2006 and Saudi's competition law of 2019, might explain some of the falling trends in some countries.



While there's no single reform that directly reduces markups across the board in the region, several initiatives can contribute to a more competitive market environment, potentially leading to lower markups. Some key channels and reforms with this potential impact include: through increasing competition, for example, trade liberalization (reducing trade barriers such as tariffs and quotas that allows for more competitive imports and puts pressure on domestic businesses to become more efficient and potentially lower prices), regulatory reforms (streamlining regulations and simplifying business registration that attract new entrants into



various sectors and fosters competition), foreign direct investment (FDI) promotion (attracting foreign companies can introduce new technologies and business practices that leads to increased competition); through enhancing market efficiency, for example, e-commerce development (encouraging e-commerce platforms that creates price transparency), antitrust regulations (strengthening antitrust laws and enforcement that prevents dominant companies from setting artificially high prices); through reforming financial sector, for example, improving access to finance for SMEs (making it easier for small and medium-sized enterprises to access loans that increases competition in the market).

It's important to note that the effectiveness of these reforms in reducing markups depends on various factors, including the specific reform implemented, the level of competition in the market, and the overall economic environment. Additionally, some reforms might have unintended consequences, so a comprehensive approach is crucial. Second, markups and market power are closely related to the price elasticity of demand: one possible reason for such trends are underlying changes in the elasticity of demand within the countries. A decrease in the price elasticity of demand means that a firm with market power can charge higher prices since demand does not adjust as much. Third, an increase in licensing agreements to build and run certain industries would also lead to an increase in market power. This can be rationalized as the government allocating incentives to diversify out of the oil sector.

Among listed firms, the construction sector has the lowest average markups, while the oil, mining, and utilities sector has the highest (Figure 7). An important reason for the low markups in construction is the absence of some large construction firms in the region from our sample. This might have reduced the weighted average markups of construction firms in our sample. Consistent with previous findings in the US, market services seem to have higher markups than other industries (De Loecker 2021). This result is robust to the use of the alternative estimation of markups as seen in Annex III.

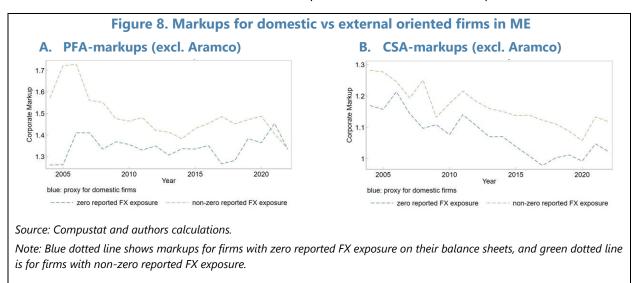
<sup>&</sup>lt;sup>9</sup> We study the effect of the introduction of the VAT on market power in Section V.c.

<sup>&</sup>lt;sup>10</sup> Indeed, through conversations with policymakers, we are aware that markups in large construction firms that are absent from our sample are high. We are also aware that this is due to multiple factors including government regulations.

<sup>&</sup>lt;sup>11</sup> Note that our definition of market services does not exactly coincide with that of De Loecker (2021) since we use an alternative industry classification code.

#### V.a.1. Who Gains from Falling Corporate Markups in ME?

The reduction in markups in the region implies a welfare gain for the customers who buy corporate goods and services from the ME. Conversely, a firm like ARAMCO charging higher markups for global (non-ME) markets implies welfare gains for the owners of ARAMCO. Who then benefits from the falling market power in ME? In this section, we plot markup evolution by firms exposed to foreign income shocks to proxy for whether the firm sells primarily to domestic consumers (Figure 8). While this is an imperfect measure in answering the question of who are the customers of ME listed firms, it is the only available marker in our data for whether a firm is foreign or not. In both of our methods, markups of firms with no FX exposure are lower than markups of firms with FX exposure. For the CSA method for estimating markups, both types of firms show a very similar reducing trend, implying that both foreign consumers and domestic consumers benefited equally from the falling trend in markups. While using the PFA method, markups for firms with no FX exposure fluctuate around a flat level, implying that domestic consumers did not benefit from the trend fall in markups. Markup for firms with FX exposure are on a downward trend implying that ME firms are becoming more competitive in global markets. Note that this result is heterogeneous by country. For example, Saudi Arabian firm markups are very similar in evolution and level between firms with FX exposure and firms without FX exposure.



#### V.a.2. Dissecting Markup Evolution in ME

Annex I reports the results of a decomposition of markup evolution into 3 terms: evolution due to net firm entry, evolution due to within firm reduction in markups, and evolution due to reallocation of economic activity. The green line in both figures reports our headline weighted markup estimators. The dashed lines present alternative scenarios starting from 2004.

The dashed red line presents a scenario where the only component of markups allowed to change is within firm markups. In both markup measures provided, the red dashed line drives the reduction in markups. This implies that, in the absence of the other changes in markups due to entry and reallocation, market power would have been lower in the region. The dashed grey line presents a scenario where the relocation of economic activity is the only term allowed to evolve whereas the rest are held constant. The relocation term stagnates in the production function approach markups and rises in the cost-share markups. For the former measure of

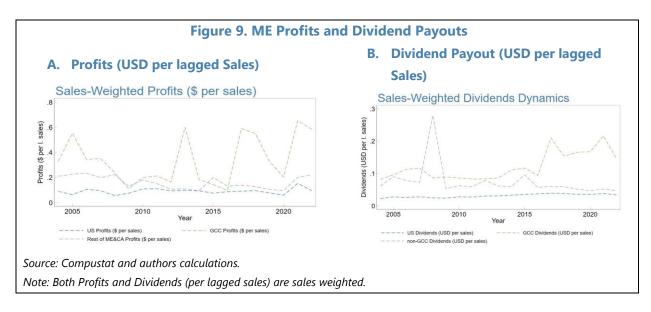
markups, this implies that reallocation did not play a part in the evolution of markups in the region. For the latter, this means reallocation of market shares to firms with higher markups led to an increase in markups. Net entry (black dashed line) remained broadly constant along the first measure of markups and fell similarly to within.

In summary, for both estimates of markups, the declining evolution is partly determined by within firm changes in markups as opposed to entry on its own. In fact, for one of the measures of markups, net firm entry plays an extremely limited role in the change of markups.

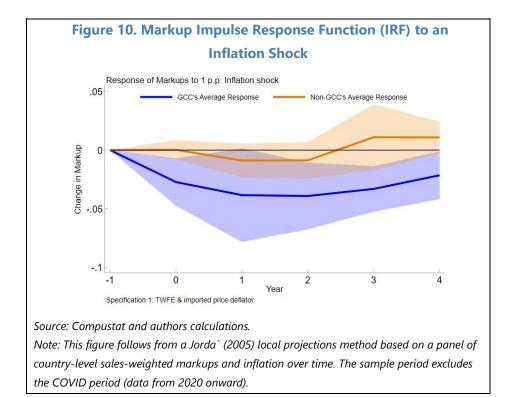
#### V.b. Markups in ME and Their Macroeconomic Implications

#### V.b.1 Markups and Inflation in ME

Our analysis suggests that a rise in global market power did not contribute to the GCC inflation surge of 2022. Corporate profits and dividend payouts in the GCC have increased robustly over 2021-2022 (Figure 9), while wages have risen relatively slowly compared to prices (IMF ME&CA REO October 2023). The rise in profits does not necessarily imply increases in monopoly power with firms deliberately raising prices above the cost of producing an additional unit of output (marginal costs). Based on our analysis of firm-level data, we have found that there has been little or no change in the markups of firms across various sectors in the ME countries (Figure 7), and the overall markups have exhibited a downward trend (except for COVID years).



There is also evidence that markups in the GCC decrease in response to a positive inflation shock, whereas in the rest of the ME markups do not change in response to inflation shocks. Figure 10 shows the results of a local projection exercise of the effect of an inflation shock on sales-weighted markups using a panel of GCC and non-GCC ME countries. The result indicates that after a 1 percent inflation shock, firms in the GCC on average reduce their markups by 0.05 units relative to an average of 1.3 after two years of the shock. Although the significance varies by country the relationship remains negative for all GCC countries. This implies that local firms absorb some of the inflationary pressures and pass less of the price changes onto consumers. However, this is not the case in the non-GCC ME sample, where firms do not seem to adjust their markups in response to an inflationary shock, indicating that they pass on price rises to their customers.

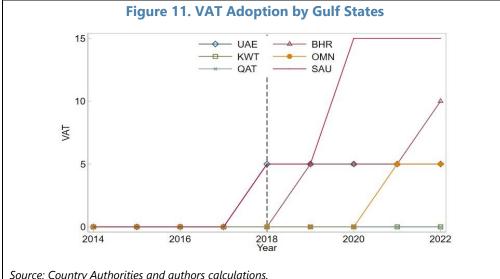


#### V.c. VAT Reform and Market Power

#### V.c.1 VAT Reform in GCC

In 2017, the GCC countries agreed a unified value-added tax (VAT) framework between the six sovereign states that make up the council. In line with unifying legislation and bolstering economic unity among its member states, the framework aimed to allow the oil dependent GCC countries to mobilize and diversify their government revenue sources. The framework stipulates a set of common legislative provisions for the application of the tax, the ability of member states to exempt certain sectors from the tax, and the registration requirements for businesses. Crucially, the treaty has two key articles. Firstly, it mandates that the signatories adopt a standard rate for VAT at 5 percent. Secondly, it requires the introduction of local laws within the signatories to facilitate the implementation of the treaty.

In practice, six years after the signing of the treaty, as shown in Figure 11, only four out of the six countries implemented the tax. In fact, Saudi Arabia in 2020 and Bahrain in 2022 hiked the rates after the initial introduction, contrary to expectations. This staggered implementation of the VAT reform and the unique regional integration initiatives allows us to study the impact of value-added taxes on firm-level market power using multiple staggered difference-in-difference identification strategies, as described in the next subsection.



Source: Country Authorities and authors calculations.

Note: VAT standard rates are shown in percentage points. GCC economies adopted the Unified VAT Framework in 2017. Kuwait and Qatar are yet to implement the tax framework of 2017. Saudi Arabia and the United Arab Emirates were the first to implement in 2018. The black dotted line indicates the year where the first set of firms were treated.

Although, the main aim of the tax reform was to diversify revenue sources and enhance revenue mobilization, here we study the unintended behavioral impact the VAT introduction had on firms market power in GCC. A long tradition in economics has noted that deviations from competitive behavior leads to welfare losses theoretically and empirically (see for example, Carson (1975), and Bilbiie et al. (2008)). Studying the ability of VAT reforms to reduce noncompetitive behavior of firms can aide policymakers in reducing welfare losses. VAT, although not the exact lever usually used to reduce market power (normally seen as the role of antitrust), can still play a role in reducing inefficiencies related to the product markets. In fact, it has been noted in the optimal tax literature that in the presence of market power, using VAT to raise revenue is optimal since it serves the dual purpose of raising revenue and reduces firm incentives to have positive markups.

#### V.c.2 Identification Strategy

We adopt multiple identification strategies to identify the effect of the VAT reform on market power in the Gulf. Those identification strategies rely on slightly different assumptions for identifying the effect.

Theoretical Model. We start by defining the behavior of a profit maximizing firm choosing its price level in the presence of a final good consumption tax and facing an inverse demand schedule Q(P). The firm optimizes its profit function by choosing its producer price as follows

$$\max_{p} P \cdot Q(p(1+\tau) - C(Q(P(1+\tau)))$$
 (5)

The first order condition implies

$$\mu_{it} =: \frac{P_{it}}{C'(Q(P(1+\tau))} = \left[1 + \frac{1}{(1+\tau)\gamma}\right]^{-1} \tag{6}$$

where  $\gamma = \frac{Q_{it}}{P \cdot Q_p}$ . Hence, the effect of a tax increase on logged firm markups is given by

$$\frac{\vartheta \ln(\mu_{it})}{\vartheta \tau} = \frac{\gamma \cdot [\gamma(1+\tau)]^{-2}}{1 + [\gamma(1+\tau)]^{-1}} \tag{7}$$

The numerator of equation 7 is negative since  $\gamma$  < 0. The sign of the denominator thus determines whether or not the tax has a correcting effect on markups. In particular, assuming demand is sufficiently elastic implies<sup>12</sup>

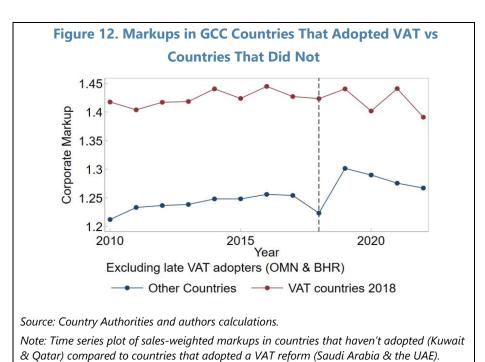
$$\frac{\vartheta \ln (\mu_{it})}{\vartheta \tau} < 0$$

This result shows that VAT reduces the wedge between output prices and marginal costs. This result was noted in previous theoretical work on optimal taxes as Delipalla & Keen (1992) noted, since intuitively the incentive to overcharge consumers is diminished as the VAT rate rises.

• Staggered Adoption Diff-in-Diff. As shown in Figure 11, VAT adoption was staggered over countries, despite all GCC states committing to the introduction of the VAT in the 2016 agreement. Our first research design relies on this staggered adoption. Specifically, we assume the following parallel trends assumption:

$$E[\Delta ln\mu_{it}^0|\tau_t = d] = E[\Delta ln\mu_{it}^0|\tau_t = 0]$$
(8)

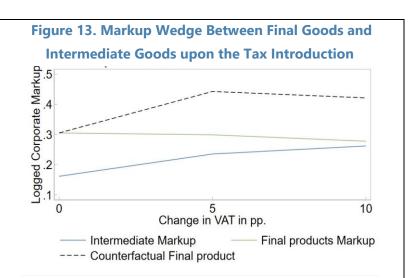
This assumption stresses that the counterfactual changes in logged markups in the absence of VAT would have been the same as changes in logged markups under no VAT implementation. We investigate robustness of our result by conducting pre-trends tests of this assumption and assuming a conditional form of this assumption



<sup>&</sup>lt;sup>12</sup> Elasticity of demand has to satisfy  $y \cdot (1 + \tau) < -1$  for revenue maximization.

based on a matched sample of firms in the treatment and control group.

Intermediate-Final Goods Diff-in-Diff. In the second exercise, we use an identification strategy that relies on a within country comparison. Firms in the intermediate sectors do not end up paying VAT and claim back their tax payments. These firms are treated as a control group for firms that pay VAT in the final sector. To define firms as intermediate or final, we use an indicator for firms operating in sectors traditionally seen as final sectors as our treatment variable.13 The identifying assumption for the treatment effect is thus that logged markups would have trended similarly in treated firms when compared to non-treated firms had the VAT reform not been implemented. This is stated in equation 9. Figure 13 presents stylized evidence of the potential impact of VAT on markups.



Source: Country Authorities and authors calculations.

Note: Non-linear plot of logged markups against changes in GCC VAT rates. Counterfactual markups in the final products sector would have been higher (black line) compared to the observed level of final products markups (green line).

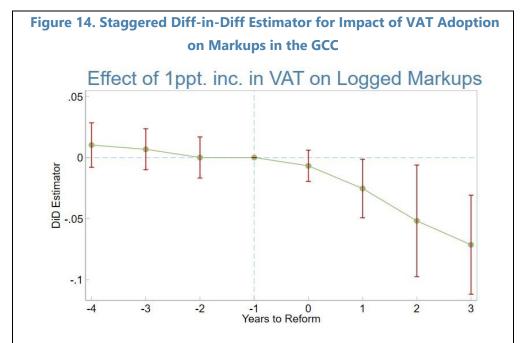
$$E[\Delta ln\mu_{it}^{0}|\tau_{t}-\tau_{t-1}=d,final] = E[\Delta ln\mu_{it}^{0}|\tau_{t}-\tau_{t-1}=d,intermediate]$$
 (9)

#### V.c.3 Identification Strategy 1: Staggered Adoption Diff-in-Diff

Based on the staggered Diff-in-Diff estimator, we find evidence of a response to the VAT reforms in the GCC states. A 1 percentage point increase in the VAT rate is associated with a 5 percent reduction in markups. The effect is not immediate, and adjustment occurs dynamically as reported in Figure 14. The figure also shows no evidence of the violation of the common trends assumption (asserted in equation 8) in the pre-reform period. We investigate multiple robustness checks around these figures. We cannot trace the effects beyond 3 years after the reform since there is not enough firms accounts after 2021 yet.

Annex IV provides evidence for internal validity of the estimates reported in Figure 14. The headline average impact reported in Column (1) of the table shows that the impact of an additional 1% point of VAT is an average reduction of 1% in markups estimated using the production function approach in the GCC sample. This result is robust to different specifications, weighting methods, and the alternative markup measure we constructed.

<sup>&</sup>lt;sup>13</sup> This includes firms operating in market services and the retail trade sector.



Source: Country Authorities and authors calculations.

Note: Staggered Diff-in-Diff event study of an increase in VAT by 1 p.p. on GCC markups. This event study relies on the estimator by de Chaisemartin & D'Haultfœuille (2020). The dependent variable is logged markups, estimated using the production function approach. We allow for firm-specific linear time trend. We control for the sales-cost of goods sold ratio, firm's sales fraction of industry, and firm's sales ratio to total sales. Standard errors are clustered at the firm level. Observations are weighted by deflated sales to study the effect of the tax policy on aggregate logged markups.

#### VI. Conclusion

Our study sheds light on corporate market power in the Middle East. We find that corporate markups in the Middle East are higher compared to the US (and Europe), for both GCC and non-GCC countries, but there has been a downward trend over the years, with the exception of the COVID-19 period. This suggests that while market power is higher in the region, it can be influenced by implemented reforms and policies.

Additionally, we observe that GCC countries have higher corporate markups compared to the rest of the ME countries. However, there has been a convergence of GCC markups towards the US and the rest of the ME

region since the onset of the COVID-19 pandemic. This suggests that external shocks and changing market dynamics can impact market power levels, even in countries with traditionally higher markups.

Furthermore, our research highlights the presence of superstar firms in the ME region, where larger firms with higher sales tend to have higher market power. This phenomenon holds true across sectors, with the oil, mining, and utilities sector exhibiting particularly high market power. Although GCC have higher markups than the rest of ME, this result is not unique to GCC. This finding underscores the importance of addressing market concentration and promoting competition to ensure a level playing field for all market participants.

Moreover, our analysis reveals that marketable services sectors and the oil, mining, and utilities industries generally have higher markups compared to other industries in the ME region. This finding highlights the need for targeted policies and regulations to ensure fair competition and prevent the concentration of market power in these sectors.

Our study finds no evidence of "greedflation" in the ME region, indicating that the relationship between market power and inflationary pressures is not significant or even negative for the GCC. This finding contributes to the understanding of the macroeconomic implications of market power in the region and suggests that other factors may be driving inflationary pressures.

Finally, our research demonstrates the potential of value-added tax (VAT) policy reforms to act as an additional deterrent of corporate market power. The introduction of VAT in some Gulf states from 2018 to 2022 has led to a reduction in market power, in addition to increasing fiscal space. This finding highlights the additional benefits of implementing VAT in promoting competition in the region.

Our study provides valuable insights into the dynamics of corporate market power in the ME region. The findings underscore the need for effective policies and regulations to address market concentration, promote competition, and ensure a level playing field for all market participants. By understanding the factors influencing market power and its implications, policymakers can design targeted interventions to foster competitive and resilient economies in the ME region.

#### References

- Acemoglu, D., Hassan, T. A. & Tahoun, A. (2017), 'The Power of the Street: Evidence from Egypt's Arab Spring', *The Review of Financial Studies* **31**(1), 1–42.

  URL: https://doi.org/10.1093/rfs/hhx086
- Ackerberg, D. A., Caves, K. & Frazer, G. (2015), 'Identification properties of recent production function estimators', *Econometrica* **83**(6), 2411–2451.
  - URL: https://onlinelibrary.wiley.com/doi/abs/10.3982/ECTA13408
- Autor, D., Dorn, D., Katz, L. F., Patterson, C. & Van Reenen, J. (2020), 'The Fall of the Labor Share and the Rise of Superstar Firms\*', *The Quarterly Journal of Economics* **135**(2), 645–709. **URL**: https://doi.org/10.1093/qje/qjaa004
- Benzarti, Y., Carloni, D., Harju, J. & Kosonen, T. (2020), 'What goes up may not come down: Asymmetric incidence of value-added taxes', *Journal of Political Economy* **128**(12), 4438–4474. **URL**: https://doi.org/10.1086/710558
- Berry, S., Levinsohn, J. & Pakes, A. (1995), 'Automobile prices in market equilibrium', *Econometrica* **63**(4), 841–90.
  - URL: https://EconPapers.repec.org/RePEc:ecm:emetrp:v:63:y:1995:i:4:p:841-90
- Bilbiie, F. O., Ghironi, F. & Melitz, M. J. (2008), Monopoly power and endogenous product variety: Distortions and remedies, Technical report, National Bureau of Economic Research.
- Bond, S., Hashemi, A., Kaplan, G. & Zoch, P. (2021), 'Some unpleasant markup arithmetic: Production function elasticities and their estimation from production data', *Journal of Monetary Economics* **121**, 1–14.
- Carson, R. (1975), 'On monopoly welfare losses: comment', *The American Economic Review* **65**(5), 1008–1014.
- Chaudhry, K. A. (1997), *The price of wealth: Economies and institutions in the Middle East*, Cornell University Press.
- de Chaisemartin, C. & D'Haultfœuille, X. (2020), 'Two-way fixed effects estimators with heterogeneous treatment effects', *American Economic Review* **110**(9), 2964–96. **URL:** https://www.aeaweb.org/articles?id=10.1257/aer.20181169
- De Loecker, J. (2021), 'Comment on (un)pleasant... by bond et al (2020)', *Journal of Monetary Economics* **121**(C), 15–18.
  - URL: https://EconPapers.repec.org/RePEc:eee:moneco:v:121:y:2021:i:c:p:15-18

- De Loecker, J., Eeckhout, J. & Unger, G. (2020), 'The Rise of Market Power and the Macroeconomic Implications\*', *The Quarterly Journal of Economics* **135**(2), 561–644. **URL:** https://doi.org/10.1093/qje/qjz041
- De Loecker, J. & Warzynski, F. (2012), 'Markups and firm-level export status', *American Economic Review* **102**(6), 2437–71.

URL: https://www.aeaweb.org/articles?id=10.1257/aer.102.6.2437

- Delipalla, S. & Keen, M. (1992), 'The comparison between ad valorem and specific taxation under imperfect competition', *Journal of Public Economics* **49**(3), 351–367. **URL:** https://www.sciencedirect.com/science/article/pii/0047272792900730
- Diez, M. F., Leigh, M. D. & Tambunlertchai, S. (2018), *Global market power and its macroeconomic implications*, International Monetary Fund.
- Diwan, I., Keefer, P. & Schiffbauer, M. (2020), 'Pyramid capitalism: Cronyism, regulation, and firm productivity in egypt', *The Review of International Organizations* **15**, 211–246.
- Hall, R. (1988), 'The relation between price and marginal cost in u.s. industry', *Journal of Political Economy* **96**(5), 921–47.

URL: https://EconPapers.repec.org/RePEc:ucp:jpolec:v:96:y:1988:i:5:p:921-47

- Hennessy, C. A. & Whited, T. M. (2005), 'Debt dynamics', *The Journal of Finance* **60**(3), 1129–1165.
  - **URL:** https://onlinelibrary.wiley.com/doi/abs/10.1111/j.1540-6261.2005.00758.x
- Jord`a, (2005), 'Estimation and inference of impulse responses by local projections', *American Economic Review* **95**(1), 161–182.

**URL:** https://www.aeaweb.org/articles?id=10.1257/0002828053828518

- Klette, T. J. & Griliches, Z. (1996), 'The inconsistency of common scale estimators when output prices are unobserved and endogenous', *Journal of Applied Econometrics* **11**(4), 343–361. **URL:** http://www.jstor.org/stable/2284929
- Loecker, J. D. & Scott, P. T. (2017), Estimating market power Evidence from the US Brewing Industry, Working Papers 17-06, Center for Economic Studies, U.S. Census Bureau. **URL:** https://ideas.repec.org/p/cen/wpaper/17-06.html
- Olley, G. S. & Pakes, A. (1996), 'The dynamics of productivity in the telecommunications equipment industry', *Econometrica* **64**(6), 1263–1297.

**URL:** http://www.jstor.org/stable/2171831

Raval, D. (2023), 'Testing the Production Approach to Markup Estimation', *The Review of Economic Studies*. rdad002.

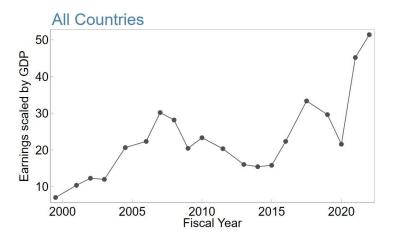
URL: https://doi.org/10.1093/restud/rdad002

Robinson, J. A. & Acemoglu, D. (2012), *Why nations fail: The origins of power, prosperity and poverty,* Profile London.

Syverson, C. (2019), 'Macroeconomics and market power: Context, implications, and open questions', *Journal of Economic Perspectives* **33**(3), 23–43. **URL:** https://www.aeaweb.org/articles?id=10.1257/jep.33.3.23

# **Annex I. Sample Representativeness**

#### A.I.a. Earnings relative to GDP

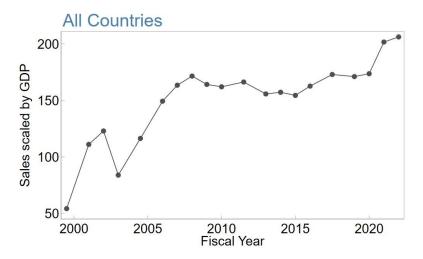


Aggregated country-level average Earnings relative to GDP.



Map of average Earnings relative to GDP in 2020.

#### A.I.b. Sales relative to GDP



Aggregated country-level average Sales relative to GDP.



Map of average Sales relative to GDP in 2020.

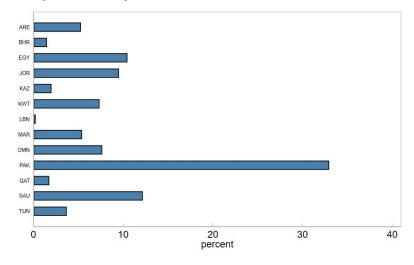
# **Annex II. Sample Description**

A.II.a. Table 1: Summary Statistics for Sample between 2004 and 2022

|     | Count of Firm-year | Sales   | COGS   | Pretax Income | PFA Markups | CSA Markups |
|-----|--------------------|---------|--------|---------------|-------------|-------------|
| ARE | 1018               | 897804  | 518217 | 149467        | 1.51        | 1.15        |
| BHR | 278                | 271782  | 173070 | 39566         | 1.42        | 1.09        |
| EGY | 2064               | 302459  | 210498 | 45091         | 1.38        | 1.13        |
| JOR | 1693               | 129969  | 103593 | 10066         | 1.12        | 1.06        |
| KAZ | 380                | 602456  | 259910 | 177650        | 2.25        | 1.66        |
| KWT | 1466               | 284078  | 181791 | 30654         | 1.62        | 1.14        |
| LBN | 46                 | 135524  | 90256  | 15971         | 1.33        | 1.03        |
| MAR | 980                | 447442  | 261570 | 69099         | 1.77        | 1.26        |
| OMN | 1372               | 152126  | 104384 | 12022         | 1.39        | 1.07        |
| PAK | 5587               | 231704  | 186559 | 21435         | 1.15        | 1.12        |
| QAT | 348                | 895081  | 530403 | 177973        | 1.58        | 1.14        |
| SAU | 2377               | 1657932 | 884909 | 577603        | 1.61        | 1.32        |
| TUN | 730                | 110941  | 80482  | 5347          | 1.16        | 1.12        |

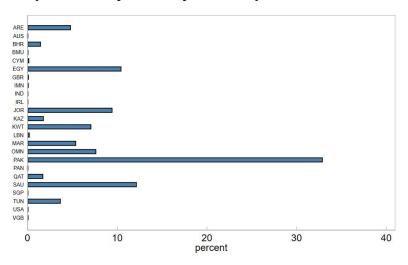
*Note.* This table reports summary statistics for the main sample of listed firms headquartered in the ME. Sales, COGS, and Pretax income are reported in '000s of 2015 USD. PFA markup estimation is explained in the text. Share estimator of markups is reported in Annex E.

A.II.b. Sample Composition: by Firm Location



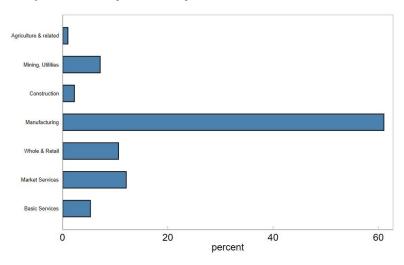
Fraction of the Sample by Location of Firm's Headquarter.

#### A.II.c. Sample Composition: by Country of Incorporation



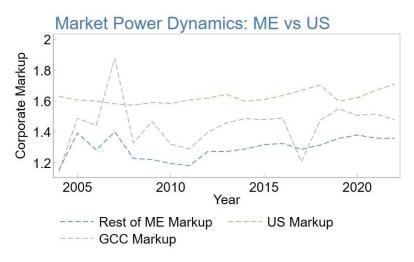
Fraction of the Sample by Firm's Country of Incorporation.

#### A.II.d. Sample Composition: by Industry

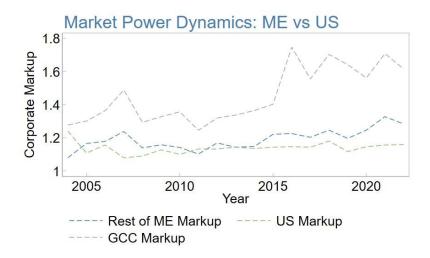


Fraction of the Sample by Firm's Industry.

#### A.II.e. PFA: excluding XSGA



Unweighted markups using production function approach excluding XSGA.



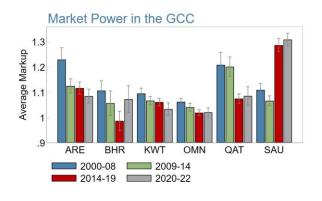
COGS-weighted markups using production function approach excluding XSGA.

# Annex III. Revenue-Variable Costs Ratio (CSA)

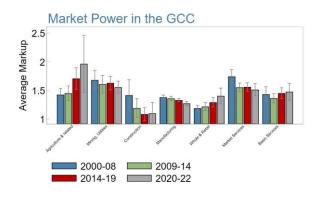


**Excluding ARAMCO** 

Markups based on CSA.

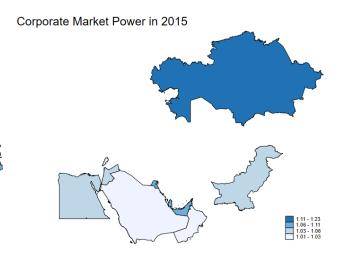


Including ARAMCO



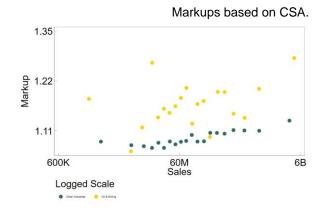
(i) GCC country

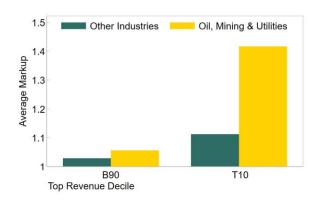
Markups based on CSA.



(ii) Industries in GCC

Weighted by COGS





(i) Superstar Phenomena

Markups based on CSA.

(ii) Deciles of Revenue

## Annex IV. Value Added Tax: Robustness Table

|              | (1)                  | (2)                      | (3)             | (4)                  | (5)                  | (6)                      |
|--------------|----------------------|--------------------------|-----------------|----------------------|----------------------|--------------------------|
| DV           | $log(\mu_{\it PFA})$ | $log(\mu_{	extit{PFA}})$ | $log(\mu_{CS})$ | $log(\mu_{\it PFA})$ | $log(\mu_{\it PFA})$ | $log(\mu_{	extit{PFA}})$ |
|              |                      |                          |                 |                      |                      |                          |
| VAT          | -0.0117**            | -0.00798**               | -0.00667***     | -0.00159**           | -0.00148**           | -0.00728**               |
|              | (0.00461)            | (0.00352)                | (0.00247)       | (0.000738)           | (0.000734)           | (0.00315)                |
| Constant     | 0.503***             | 0.382***                 | 0.263***        | -0.176***            | -0.134**             | 18.11                    |
|              | (0.0280)             | (0.0206)                 | (0.0150)        | (0.0636)             | (0.0642)             | (11.62)                  |
| Observations | 3,691                | 3,691                    | 3,606           | 3,691                | 3,691                | 3,691                    |
| R-squared    | 0.809                | 0.870                    | 0.909           | 0.983                | 0.984                | 0.818                    |
| Firm FE      | Υ                    | Υ                        | Υ               | Υ                    | Υ                    | Υ                        |
| Year Trend   | N                    | N                        | N               | N                    | N                    | Υ                        |
| Controls     | N                    | N                        | N               | Υ                    | Υ                    | N                        |
| Weight       | Sales                | COGS                     | Sales           | Sales                | Sales                | Sales                    |
| GCC Sample   | Υ                    | Υ                        | Υ               | Υ                    | Υ                    | Υ                        |
| Oil          | N                    | N                        | N               | N                    | Υ                    | Υ                        |
|              |                      |                          |                 |                      |                      |                          |

Note. This table reports robustness of main VAT result to different controls. The table reports the effect of a 1% point increase in value-added tax on market power. The table focuses on the sample of GCC headquartered firms. PFA markup estimation is explained in the text. Share estimator markups is explained in Annex E. Clustered standard errors in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

# **Annex V. Country Specific Results**

#### A.VI.a. Saudi Arabia

Figure 17: Markups using PFA in Saudi Arabia.

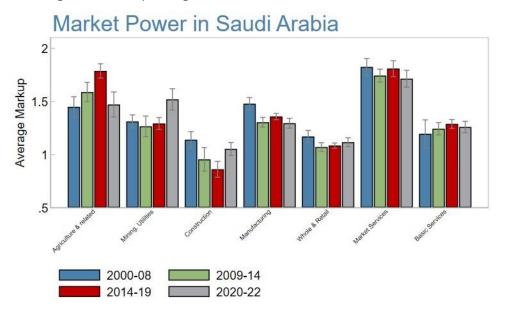
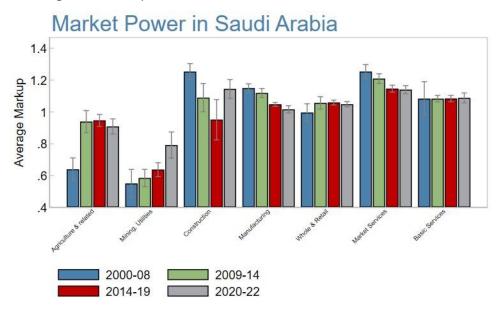


Figure 18: Markups based on CSA in Saudi Arabia.



#### A.VI.b. Egypt

Figure 19: Markups using PFA in Egypt.

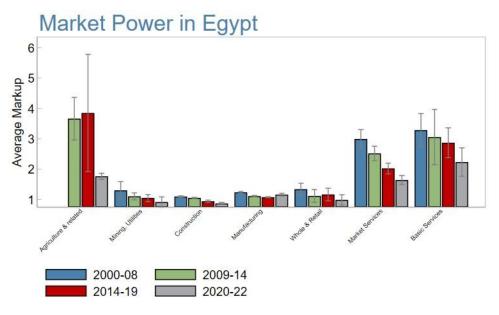
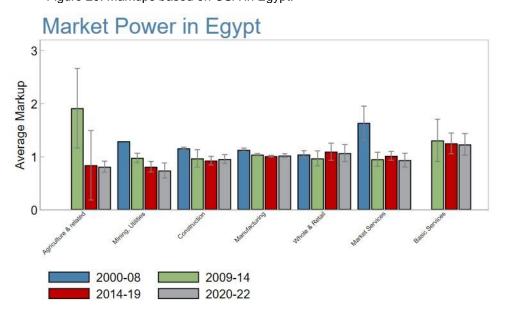
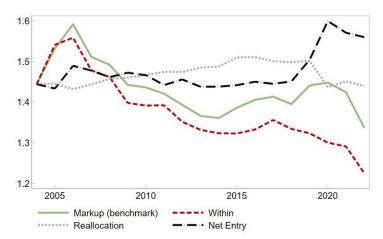


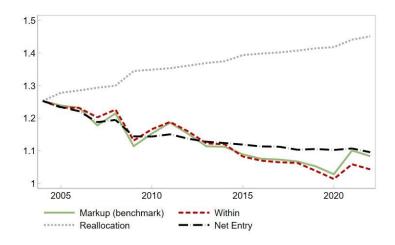
Figure 20: Markups based on CSA in Egypt.



# **Annex VI. Breakdown of Markups Dynamics**

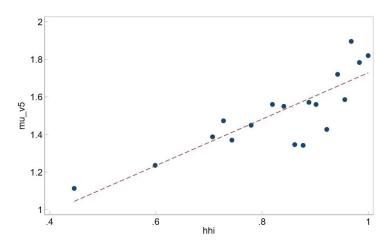


PFA markups excluding ARAMCO.

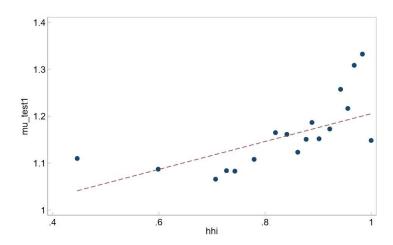


CSA-markups excluding ARAMCO.

# Annex VII. Correlation between Markup Measures and HH-index Measure of Concentration

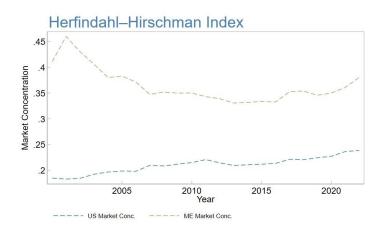


PFA markups excluding ARAMCO.

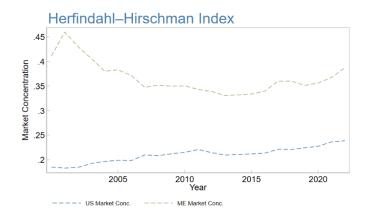


CSA-markups excluding ARAMCO.

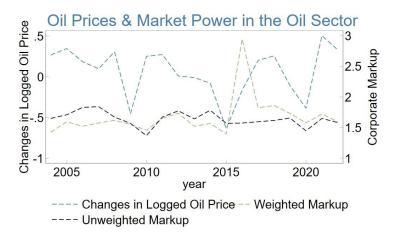
# **Annex IVIII. Additional Results**



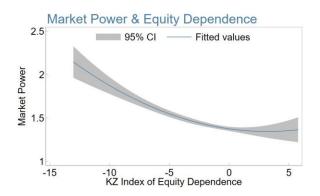
HH-index based on 4-digit industry classification excluding ARAMCO.

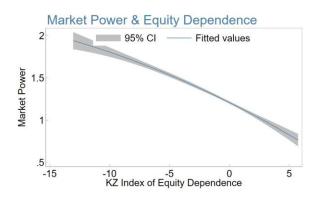


HH-index based on 4-digit industry classification including ARAMCO.



Correlation between unweighted markups & price is 0.23; weighted markups & price is -0.04  $\,$ 





(i) Gulf Cooperation Council

(ii) Rest of Middle East

KZ-index of equity dependence is negatively correlated with PFA Markups.

